Supplementary Information for

"Electricity demand has not become more price-responsive despite ninety years of technological change

Peter Kudela[1], Tomas Havranek[1,3,4], Zuzana Irsova[1,2,3],
Anna Kudelova[1], and Vojtech Sikl[1]

[1]Charles University, Prague
[2]Anglo-American University, Prague
[3]Meta-Research Innovation Center, Stanford
[4]CEPR, London

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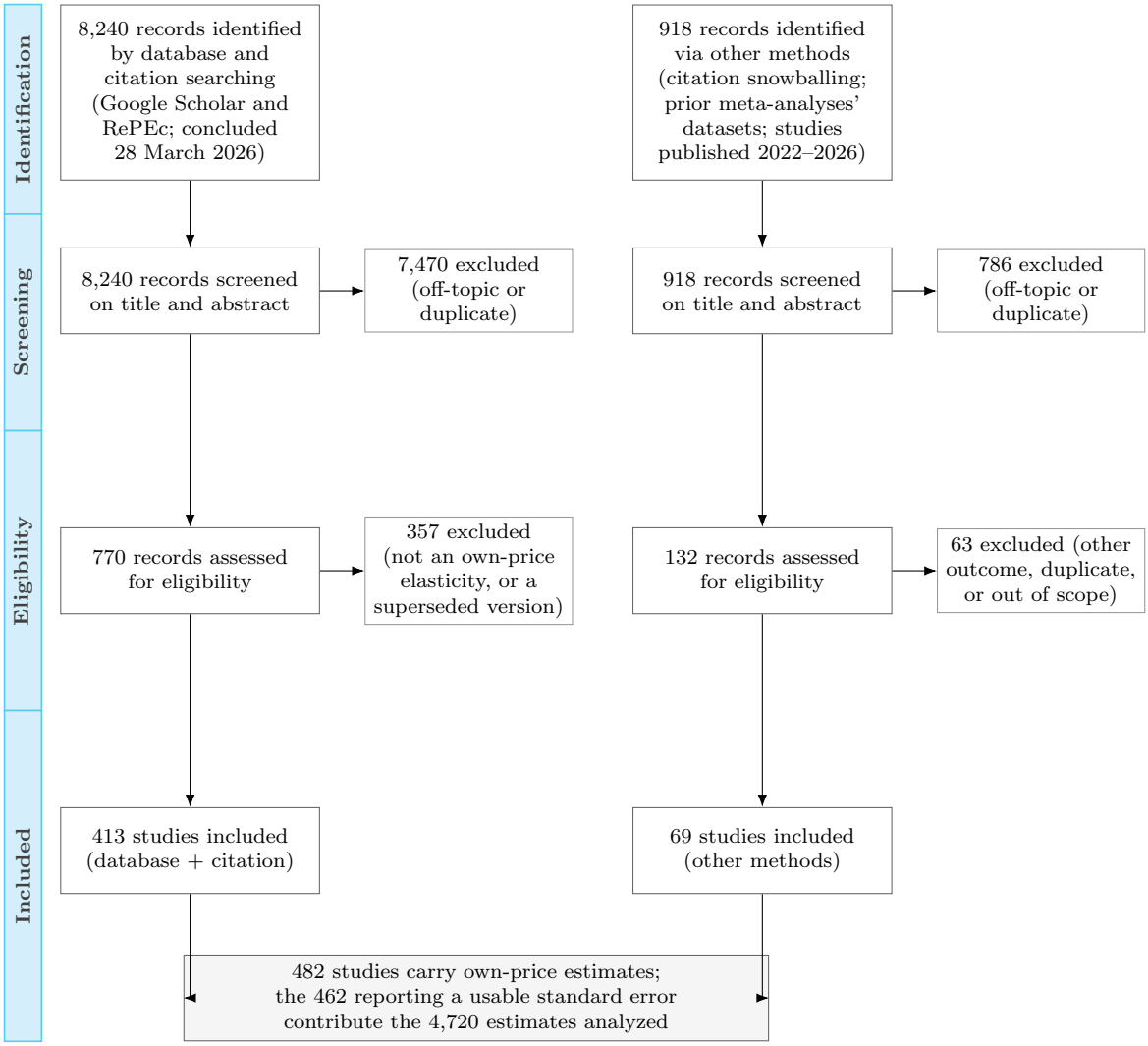
Appendix S1 Literature search and summary statistics

The corpus was assembled by systematic database search and citation tracking; Figure S1 reports the flow. The database-and-citation arm builds on a systematic Google Scholar and RePEc search concluded on 28 March 2026, which produced a 413-study base collection; the other-methods arm adds backward and forward citation snowballing, the datasets of prior electricity-demand meta-analyses, and studies published through early 2026, extending the collection to 482 studies. Reporting follows the meta-analysis guidelines of Havranek *et al.* (2020); for the use of AI, we also follow the guidance of Cook *et al.* (2026a) and Cook *et al.* (2026b). To qualify for the analysis, a study had to report an own-price elasticity of electricity demand, not a cross-price or income elasticity, together with a usable measure of uncertainty: a standard error, a t-statistic, or a p-value from which a standard error could be recovered. The 462 studies meeting this rule contribute the 4,720 own-price elasticity estimates that make up the corpus; the remaining 20 studies report no usable measure of uncertainty and are excluded. We convert the 745 compensated estimates that report both an income elasticity and a usable standard error; together with the 2,579 directly reported Marshallian estimates they make up the 3,324-estimate baseline. The 1,396 estimates held out of that baseline are of three kinds: 859 (from 77 studies) that cannot be placed on the Marshallian footing for want of a usable income elasticity to convert, 420 whose elasticity is indirectly derived from an inverted formula, and 117 whose standard error is imputed from a reported significance level (all from a single study).

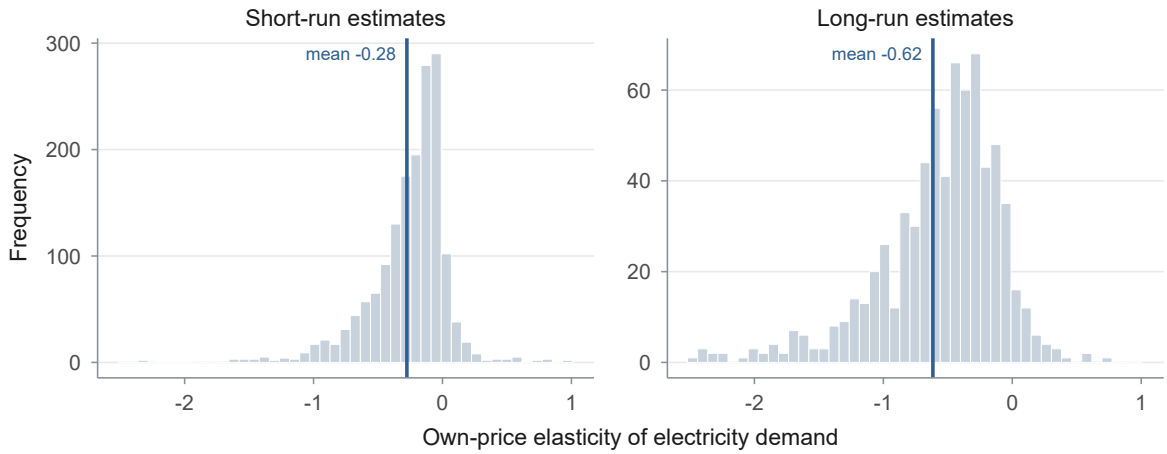
Figure S2 plots the distribution of the estimates, and Table S1 and Table S2 summarize them across the main subsamples for the short and long run separately. The estimates are

Figure S1: PRISMA flow diagram of the literature search



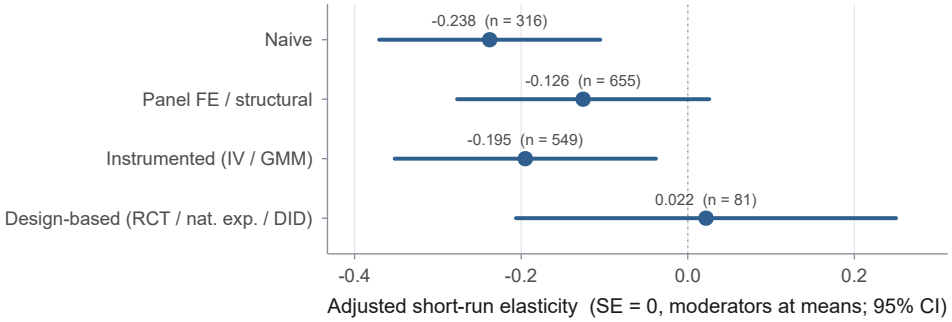
Notes: Two-stream PRISMA flow for the electricity-demand price-elasticity corpus; reporting follows the meta-analysis guidelines of Havranek *et al.* (2020) and, for AI use, Cook *et al.* (2026a) and Cook *et al.* (2026b). *Database and citation arm:* a systematic Google Scholar and RePEc search (own-price elasticity of electricity demand paired with residential, industrial, tariff, and estimation-method terms) with forward and backward citation searching, concluded 28 March 2026, retrieves 8,240 records; 7,470 are removed at de-duplication and title/abstract screening and 770 are assessed in full text, of which 357 are excluded (a cross-price or income elasticity rather than own-price, or a superseded or out-of-scope version), leaving 413 studies. *Other-methods arm:* backward and forward citation snowballing on the most-cited estimates, the datasets of prior electricity-demand meta-analyses (Dahl, 2011; Espey & Espey, 2004; Labandeira *et al.*, 2017; Zhu *et al.*, 2018; Zabaloy & Viego, 2022; Fatima, 2023), and studies published between 2022 and 2026 identify 918 further records; 132 are assessed and 69 are included. Together the two arms give $413 + 69 = 482$ studies with own-price estimates; applying the inclusion rule (a usable standard error, t-statistic, or p-value) leaves 462 studies contributing 4,720 estimates. PRISMA = Preferred Reporting Items for Systematic Reviews and Meta-Analyses (Page *et al.*, 2021). Data, code, and additional materials are available in an online appendix at `https://meta-analysis.cz/electricity`.

Figure S2: Distribution of the price-elasticity estimates



Notes: Histograms of the own-price elasticity of electricity demand (Marshallian-equivalent, reported values), for the 1,647 short-run and 723 long-run headline estimates (usable standard error); the two panels share a common horizontal scale, truncated to $[-2.5, 1]$ so that 8 short-run and 13 long-run outlying estimates fall outside the frame (trimmed from the plot only, and kept in every statistic). The vertical line marks the simple (unweighted) mean of each subsample, printed alongside. The long-run estimates sit farther from zero (more elastic) than the short-run estimates, reproducing the horizon gradient of Section 5.

Figure S3: The adjusted ladder (the key result of Section 3)



Notes: Graphical version of Table 1: adjusted short-run elasticity by tier with 95% study-clustered intervals.

overwhelmingly negative, as demand theory requires; wrong-signed (positive) estimates are rare, the mass concentrating in a narrow band between roughly -0.5 and 0 in the short run and shifting away from zero at the long-run horizon.

Table S3 collects the statements of the rising-flexibility premise that this paper tests, in the words of those who hold it: five sources across four decades, from the earliest time-of-use pricing experiments to the field's most recent quantitative synthesis, none of which estimates a time trend in price responsiveness.

Table S1: Summary of the short-run estimates by subsample

			Unweighted		Weighted	
Subsample	N	Studies	Mean	95% CI	Mean	95% CI
All estimates	1,647	226	−0.27	[−0.28, −0.25]	−0.11	[−0.11, −0.11]
Identification tier						
Design-based (tiers 1–3)	81	8	−0.09	[−0.15, −0.03]	−0.08	[−0.08, −0.07]
Instrumented (IV)	549	52	−0.30	[−0.33, −0.28]	−0.10	[−0.11, −0.10]
Panel FE / structural	655	98	−0.21	[−0.23, −0.18]	−0.15	[−0.15, −0.15]
Naive	316	70	−0.37	[−0.41, −0.32]	−0.11	[−0.11, −0.11]
Demand sector						
Residential	1,048	165	−0.29	[−0.31, −0.27]	−0.18	[−0.18, −0.18]
Industrial	220	56	−0.34	[−0.40, −0.27]	−0.07	[−0.08, −0.07]
Commercial	162	43	−0.24	[−0.29, −0.19]	−0.09	[−0.09, −0.08]
Data structure						
Time-series	555	104	−0.21	[−0.24, −0.18]	−0.04	[−0.05, −0.04]
Cross-section	128	26	−0.48	[−0.55, −0.41]	−0.26	[−0.26, −0.25]
Panel	964	106	−0.27	[−0.29, −0.25]	−0.13	[−0.13, −0.13]
Price measurement						
Average price	1,030	166	−0.26	[−0.28, −0.25]	−0.11	[−0.11, −0.11]
Marginal price	401	48	−0.30	[−0.34, −0.26]	−0.05	[−0.06, −0.05]
Tariff regime						
Time-of-use	210	19	−0.17	[−0.22, −0.12]	−0.01	[−0.01, −0.01]
Increasing-block	405	40	−0.31	[−0.34, −0.29]	−0.16	[−0.16, −0.16]
Decreasing-block	118	23	−0.27	[−0.33, −0.22]	−0.10	[−0.10, −0.09]
Geographic aggregation						
Country-level	396	87	−0.22	[−0.25, −0.18]	−0.08	[−0.08, −0.07]
Sub-country level	1,244	142	−0.28	[−0.30, −0.27]	−0.14	[−0.14, −0.14]
Region						
United States	634	70	−0.25	[−0.27, −0.22]	−0.09	[−0.10, −0.09]
Europe	273	37	−0.26	[−0.30, −0.23]	−0.08	[−0.08, −0.08]
Other	781	127	−0.30	[−0.32, −0.27]	−0.17	[−0.17, −0.17]

Notes: Headline short-run sample (Marshallian-equivalent, usable standard error; 1,647 estimates from 226 studies). *Unweighted* is the simple mean of the 1%-winsorized elasticity with its analytic 95% interval ($\pm 1.96\,\text{s.d.}/\sqrt{N}$); *Weighted* is the inverse-variance (precision-)weighted mean with its fixed-effect 95% interval. Subsamples are defined by indicator variables and are not mutually exclusive, so they need not partition the full sample (an estimate can be, for example, both residential and panel). Tiers 1–3 are pooled as design-based; "Sub-country level" pools regional, city, and micro-level estimates.

Table S2: Summary of the long-run estimates by subsample

			Unweighted		Weighted	
Subsample	N	Studies	Mean	95% CI	Mean	95% CI
All estimates	723	151	−0.58	[−0.62, −0.55]	−0.38	[−0.38, −0.38]
Identification tier						
Design-based (tiers 1–3)	4	1	−0.32	[−0.36, −0.28]	−0.29	[−0.35, −0.23]
Instrumented (IV)	190	36	−0.68	[−0.76, −0.61]	−0.24	[−0.25, −0.23]
Panel FE / structural	397	81	−0.55	[−0.60, −0.50]	−0.25	[−0.26, −0.25]
Naive	121	36	−0.54	[−0.62, −0.47]	−0.50	[−0.50, −0.50]
Demand sector						
Residential	464	106	−0.59	[−0.64, −0.55]	−0.43	[−0.43, −0.43]
Industrial	114	43	−0.50	[−0.60, −0.40]	−0.17	[−0.19, −0.16]
Commercial	73	29	−0.57	[−0.70, −0.44]	−0.28	[−0.31, −0.26]
Data structure						
Time-series	262	80	−0.56	[−0.64, −0.49]	−0.09	[−0.10, −0.09]
Cross-section	122	16	−0.64	[−0.70, −0.58]	−0.49	[−0.49, −0.48]
Panel	339	59	−0.58	[−0.63, −0.53]	−0.31	[−0.31, −0.30]
Price measurement						
Average price	587	126	−0.57	[−0.61, −0.53]	−0.42	[−0.42, −0.42]
Marginal price	52	17	−0.70	[−0.83, −0.57]	−0.49	[−0.50, −0.47]
Tariff regime						
Time-of-use	30	6	−1.00	[−1.27, −0.72]	−0.29	[−0.32, −0.26]
Increasing-block	152	23	−0.53	[−0.60, −0.46]	−0.47	[−0.48, −0.47]
Decreasing-block	45	10	−0.79	[−0.92, −0.65]	−0.85	[−0.87, −0.84]
Geographic aggregation						
Country-level	265	71	−0.52	[−0.59, −0.46]	−0.40	[−0.40, −0.39]
Sub-country level	449	80	−0.62	[−0.67, −0.58]	−0.37	[−0.37, −0.37]
Region						
United States	177	30	−0.57	[−0.63, −0.51]	−0.54	[−0.55, −0.53]
Europe	200	38	−0.58	[−0.66, −0.50]	−0.27	[−0.28, −0.27]
Other	353	94	−0.60	[−0.65, −0.54]	−0.41	[−0.41, −0.41]

Notes: Headline long-run sample (Marshallian-equivalent, usable standard error; 723 estimates from 151 studies). Columns and conventions are as in Table S1; a subsample present at the short but not the long horizon is omitted.

Table S3: The rising-flexibility premise, 1985–2025

Source	The premise, in the source's words
Aigner (1985, p. 39)	"[A]ppliance choices will be made with an eye to TOU response; new appliances will become widely available… the pricing strategy surely must be even more desirable in the long-run."
Faruqui & Sergici (2010)	Roughly 40% of the $40 billion advanced-metering investment "could be covered by reductions in power generation costs that could be brought about through demand response"; dynamic pricing "will have to be instituted."
Torriti (2014)	"[M]uch will change with the introduction of In Home Displays (or 'smart metres') in every home."
Huntington *et al.* (2019)	"[I]mportant transitions may eventually reshape future energy demand responses."
Kahn-Lang *et al.* (2025)	Time-based rates "may be a useful and efficient tool for encouraging demand-side flexibility"; the response "may be increasingly important as the power sector becomes more decarbonized."

Notes: Verbatim quotations from the cited works. The statements span four decades and include the field's most recent quantitative synthesis, which treats the future path of responsiveness as an open prediction problem. The Aigner (1985) entry records the profession's prevailing expectation at the time rather than his own belief; he himself doubted that the effect would grow. None of these sources estimates a time trend in price responsiveness.

Appendix S2 Studies included in the meta-analysis

Table S4 lists every study contributing estimates to the corpus. Studies are ordered alphabetically by first author; each entry links by hyperlink to its full reference in the bibliography.

Table S4: Studies included in the meta-analysis

Ackah *et al.* (2014)	Egorova & Volchkova (2004)	Mamkhezri (2025)
Acton *et al.* (1976)	Ekpo *et al.* (2011)	Mansur *et al.* (2005)
Acton *et al.* (1980)	El-Shazly (2013)	Arthur *et al.* (2012)
ADNDE (1981)	Eltony & Mohammad (1993)	Martins *et al.* (2023)
Adom (2017)	Eltony (1995)	Masike & Vermeulen (2022)
Agostini *et al.* (2015)	Eltony & Hoque (1996)	Matsui (1979)
Akmal & Stern (2001)	Eltony & Hajeeh (1999)	Matsukawa *et al.* (1993)
Al-Faris (2002)	Eltony (2004)	Mattos & Lima (2005)
Al-Sahlawi (1999)	Eltony (2006)	McFadden *et al.* (1977)
Al-Bajjali & Shamayleh (2018)	Eltony & Al-Awadhi (2007)	Mchugh (1977)
Al Irsyad *et al.* (2018)	Enrich *et al.* (2024)	McRae & Meeks (2016)
Alberini *et al.* (2011)	Erickson *et al.* (1973)	Meher (2020)
Alberini & Filippini (2011)	Eskeland *et al.* (1994)	Mendoza & Vargas (1987)
Alberini *et al.* (2019)	Eskeland & Mideksa (2010)	Mikayilov *et al.* (2017)
Alberini *et al.* (2022)	Fan *et al.* (2019)	Mikayilov *et al.* (2020)
Alter & Syed (2011)	Fatai *et al.* (2003)	Miller & Alberini (2016)
Amarawickrama & Hunt (2008)	Fell *et al.* (2014)	Modiano (1984)
Amusa *et al.* (2009)	Filippini (1995a)	Moghaddam (2003)
Anderson (1971)	Filippini (1995b)	Morovat *et al.* (2019)
Anderson (1973a)	Filippini (1999)	Mount *et al.* (1974)
Anderson (1973b)	Filippini & Pachauri (2004)	Mount & Chapman (1979)
Anderson (1974)	Filippini (2011)	Mountain & Hsiao (1989)
Andrade & Lobao (1997)	Filippini *et al.* (2018)	Mountain (1989)
Andrikopoulos *et al.* (1989)	Fisher & Kaysen (1962)	Mountain *et al.* (1989)
Ang *et al.* (1992)	Fouquet (1995)	Munley *et al.* (1990)
Apte (1983)	Frondel *et al.* (2019)	Murray *et al.* (1978)
Archibald *et al.* (1982)	Fullerton *et al.* (2012)	Nagata (2001)
Arisoy & Ozturk (2014)	Gabreyohannes (2010)	Nakajima & Hamori (2010)
Aroonruengsawat *et al.* (2012)	Gam & Rejeb (2012)	Nakajima (2010)
Arsenault *et al.* (1995)	Garbacz (1983a)	Narayan & Smyth (2005)
Asadoorian *et al.* (2006)	Garbacz (1983b)	Narayan *et al.* (2007)
Asadoorian *et al.* (2008)	Garbacz (1983c)	Nasir *et al.* (2008)
Aslam & Ahmad (2023)	Garbacz (1984a)	Okajima & Okajima (2013)
Atakhanova & Howie (2005)	Garbacz (1984b)	Oliveira (1993)
Atalla & Hunt (2016)	Garbacz (1984c)	Olivia & Gibson (2008)
Athukorala & Wilson (2010)	Garbacz (1986)	Ortiz-Velazquez *et al.* (2017)
Athukorala *et al.* (2019)	Garcia-Cerrutti (2000)	Ota *et al.* (2018)
Atkinson (1979a)	Gautam & Paudel (2018)	Otero Prada (1984)
Atkinson (1979b)	Ghaderi *et al.* (2006)	Otsuka & Haruna (2016)
Auray *et al.* (2019)	Gill & Maddala (1978)	Otsuka (2017)
Azevedo *et al.* (2011)	Glakpe & Fazzolare (1985)	Otsuka (2023)
Babatunde & Shuaibu (2009)	Gollnick (1975)	Parfomak & Lave (1996)
Badri (1992)	Goo *et al.* (2019)	Parhizgari & Davis (1978)
Balabanoff (1994)	Green *et al.* (1986)	Park & Acton (1984)
Banda & Verdugo (2007)	Gundimeda & Kohlin (2008)	Parti & Parti (1980)
Barnes *et al.* (1981)	Guo & Tybout (1994)	Paul *et al.* (2009)
Barrientos *et al.* (2018)	Halicioglu (2007)	Pellini (2021)
Basu (1976)	Hall (1986)	Pesaran *et al.* (1999)
Baughman *et al.* (1979)	Halvorsen (1975)	Pielow *et al.* (2012)
Beenstock *et al.* (1999)	Halvorsen (1976)	Pindyck (1979)
Beierlein *et al.* (1981)	Halvorsen (1977)	Pindyck (1980)
Bekhet & Othman (2011)	Halvorsen & Ford (1979)	Pitt (1985)
Belanger *et al.* (1990)	Halvorsen & Larsen (2001)	Polemis (2007)
Benavente *et al.* (2005)	Ham *et al.* (1997)	Polemis & Dagoumas (2013)
Bernard *et al.* (1987)	Hartman & Werth (1981)	Pourazarm & Cooray (2013)
Bernard *et al.* (1996)	Hasanov *et al.* (2016)	Qi *et al.* (2008)
Bernard *et al.* (2011)	Hausman (1979)	Rahman (1982)
Berndt *et al.* (1980)	Hawkins (1975)	Rai *et al.* (2014)
Berndt & Samaniego (1984)	Hawkins (1977)	Ramcharran (1988)
Bernstein & Griffin (2006)	Hawkins (1978)	Ramcharran (1990)

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Table S4 continued

Bernstein & Madlener (2011)
Bernstein & Madlener (2015)
Betancourt (1981)
Bianco *et al.* (2009)
Bianco *et al.* (2010)
Bigano *et al.* (2006)
Bigerna & Bollino (2014)
Bilgili *et al.* (2010)
Biswas *et al.* (2026)
Bjerkholt & Rinde (1983)
Bjorner *et al.* (2001)
Blasch *et al.* (2017)
Blazquez *et al.* (2013)
Blundell & Robin (1999)
Boogen *et al.* (2017)
Borenstein (2009)
Bose & Shukla (1999)
Botero *et al.* (1990)
Branch (1993)
Brenton (1997)
Burke & Kurniawati (2018)
Burke & Abayasekara (2018)
Bushnell & Mansur (2005)
Byrne *et al.* (2021)
Cabral *et al.* (2020)
Campbell (2018)
Cao *et al.* (2019)
Cao *et al.* (2023)
Cargill & Meyer (1971)
Carlevaro & Spierer (1983)
Casarin & Delfino (2011)
Cavoulacos & Caramanis (1983)
Cebula & Herder (2009)
Cebula (2012)
Cetinkaya & Bagdadioglu (2015)
Chama (2012)
Chang & Chern (1981a)
Chang & Chern (1981b)
Chang & Hsing (1991)
Chang & Martinez-Chombo (2003)
Chang *et al.* (2014)
Chaudhary *et al.* (1999)
Chaudhry (2010)
Chen & Wang (2025)
Chern (1975)
Chern (1978)
Chern & Bouis (1988)
Chindarkar & Goyal (2019)
Chishti (1993)
Choi (2002)
Christodoulakis & Kalyvitis (1997)
Christopoulos (2000)
Chung & Aigner (1981)
Cialani & Mortazavi (2018)
Cicchetti & Smith (1975)
CISEPA (1998)
Considine (2000)
Costa & Kahn (2010)
Costa & Kahn (2011)
Coughlin (1995)
Cuddington & Dagher (2015)
Dahan (1996)
Damien *et al.* (2019)
Davis (2008)
De Cian *et al.* (2007)
Delfino & Givogri (1979)
Delfino (1995)
Denton *et al.* (1999)
He & Lambert (2004)
Henderson (1983)
Henriksson *et al.* (2014)
Henson (1984)
Herriges & King (1994)
Hesse & Tarkka (1986)
Hieronymus (1976)
Hill *et al.* (1983)
Hirth *et al.* (2022)
Hogan (1989)
Holtedahl & Joutz (2004)
Hondroyiannis (2004)
Horowitz (2007)
Hosoe & Akiyama (2009)
Houston (1982)
Houthakker (1951)
Houthakker *et al.* (1974)
Houthakker (1980)
Hsiao *et al.* (1989)
Hsiao & Mountain (1994)
Hsing (1994)
Hsueh & Gerner (1986)
Hu *et al.* (2019)
Huang *et al.* (2024)
Hughes-Cromwick (1985)
Hung & Huang (2015)
Huntington & Soffer (1982)
Hyndman *et al.* (1980)
Idso *et al.* (2024)
IEEJ (1986)
Ilmakunnas & Torma (1989)
Inglesi-Lotz & Blignaut (2011)
Inglesi-Lotz (2011)
Inglesi (2010)
Iqbal (1986)
Irffi *et al.* (2006)
Ishaque (2018)
Ito (2014)
Ito *et al.* (2018)
Jaffee *et al.* (1982)
Jamil & Ahmad (2011)
Javid & Qayyum (2014)
Jeong *et al.* (2011)
Jeong & Kim (2025)
Jessoe & Rapson (2014)
Jia *et al.* (2021)
Jin & Kim (2022)
Johnsen (2001)
Jones (1995)
Jungeilges & Dahl (1986)
Kamerschen & Porter (2004)
Karbuz *et al.* (1997)
Kaserman & Mayo (1985)
KEEI (1989)
Keng (1991)
Khan & Qayyum (2009)
Khan & Abbas (2016)
Khanna *et al.* (2016)
Khazzoom (1986)
Knaut & Paulus (2016)
Kohler & Mitchell (1984)
Kohler (2014)
Kokkelenberg & Mount (1993)
Kolstad & Lee (1993)
Krishnamurthy & Kristrom (2015)
Kwon *et al.* (2016)
Lanot & Vesterberg (2021)
Lareau & Darmstadter (1982)
Rapson (2014)
Romero-Jordan *et al.* (2014)
Rosas-Flores (2017)
Rossi & Tansini (1989)
Roth (1981)
Rouhani *et al.* (2022)
Roy (1986)
Ryan *et al.* (1996)
Ryu *et al.* (2021)
Sa'ad (2009)
Sabir *et al.* (2013)
Saddler *et al.* (1980)
Saddler & Donnelly (1983)
Saha & Bhattacharya (2018)
Salari & Javid (2016)
Saunoris & Sheridan (2013)
Schulte & Heindl (2017)
Schwarz (1984)
Setshedi & Mah (2025)
Shaffer (2020)
Shi *et al.* (2012)
Shin (1981)
Shin (1985)
Silk & Joutz (1997)
Silva *et al.* (2018)
Siqueira *et al.* (2006)
Smith (1980)
Solaymani *et al.* (2015)
Sterner (1985)
Sterner (1989)
Su (2018)
Sun & Lin (2013)
Sutherland (1983a)
Sutherland (1983b)
S&Z Consultores (1999)
Talbi *et al.* (2022)
Tambe & Joshi (2014)
Tanishita (2019)
Tatli (2017)
Taylor *et al.* (1977)
Taylor (1979)
Taylor *et al.* (2005)
Terza (1986)
THEC (1983)
Tiwari (2000)
Tiwari & Menegaki (2019)
Tran *et al.* (2023)
Tserkezos (1992)
Turkekul & Unakitan (2011)
Twerefou & Abeney (2020)
Uhr *et al.* (2019)
Urga & Walters (2003)
Uri (1977)
Uri (1978)
Uri (1979b)
Uri (1979c)
Uri (1979d)
Uri (1979e)
Uri (1982)
Uri (1983)
Vashist (1984)
Veall (1983)
Veall (1987)
Velez *et al.* (1987)
Verleger (1973)
Vlachou & Samouilidis (1986)
Volland & Tilov (2018)
Wakashiro (2019)

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Table S4 continued

Denton *et al.* (2003)	Larsson (2004)	Walfridson (1987)
Dergiades & Tsoulfidis (2008)	Larsson (2006)	Walker (1979)
Deryugina *et al.* (2020)	Laumas & Williams (1981)	Wang (1985)
De Vita *et al.* (2006)	Vasquez Lavin *et al.* (2011)	Wang & Mogi (2017)
Diabli (1998)	Lee & Lee (2010)	Wang *et al.* (2020)
Dicembrino & Trovato (2013)	Lee & Chiu (2011)	Wang & Lin (2021)
Dilaver & Hunt (2011a)	Li *et al.* (2020)	Westley (1984a)
Dilaver & Hunt (2011b)	Liddle & Huntington (2021)	Westley (1984b)
Dobozi (1988)	Lijesen (2007)	Westley (1989a)
Dodgson *et al.* (1990)	Lillard & Acton (1981)	Westley (1989b)
Donatos & Mergos (1989)	Lim *et al.* (2014)	Westley (1992)
Donatos & Mergos (1991)	Lin & Ouyang (2014)	Wijemanne (1987)
Dong & Kim (2018)	Lin & Zhu (2020)	Wilson (1974)
Dong *et al.* (2020)	Liu (2005)	Woo *et al.* (2018)
Donnelly (1984)	Lohani (1992)	Yang (1978)
Donnelly & Diesendorf (1984)	Loi & Ng (2018)	Yang & Liang (2023)
Donnelly & Saddler (1984)	Lyman (1994)	Yin *et al.* (2016)
Donnelly (1985)	Lynk (1989)	Yoo *et al.* (2007)
Donnelly (1987)	Ma & Stern (2016)	Young *et al.* (1983)
Douthitt (1989)	Macroconsult (2001)	Yu & Xin (2020)
Dubin (1985)	Maddala *et al.* (1994)	Zachariadis & Pashourtidou (2007)
Dulleck & Kaufmann (2004)	Maddala *et al.* (1997)	Zarnikau (1990)
Duncan & Binswanger (1976)	Maddigan *et al.* (1983)	Zhai *et al.* (2023)
Dunstan & Schmidt (1988)	Maddock & Castan (1991)	Zhou & Teng (2013)
Durant (1990)	Madlener (2011)	Ziramba (2008)
Durmaz *et al.* (2020)	Mahmud & Chishti (1990)	Ziramba & Kavezeri (2012)

Notes: The 462 studies that make up the meta-analysis corpus, ordered alphabetically by first author and reading down each column. Each entry links to the reference list by hyperlink.

Appendix S3 Coding the identification ladder

Each estimate's price-identification strategy is coded from the primary study's description of its price variation, not from its estimator label alone. Tier 1 (randomized): experimental assignment of tariffs or randomized encouragement (e.g., Wolak, 2011; Jessoe & Rapson, 2014; Ito *et al.*, 2018). Tier 2 (natural experiment): plausibly exogenous price variation from a mandated change or a spatial boundary design (e.g., Ito, 2014, exploiting spatial discontinuities across utility service-area boundaries). Tier 3: difference-in-differences designs against untreated comparison groups (e.g., Deryugina *et al.*, 2020, which tracks municipally-aggregated communities against untreated neighbors). Tier 4: instrumented specifications (instrumental variables, IV; two- and three-stage least squares, 2SLS and 3SLS; and the generalized method of moments, GMM), including supply-shifter and Lewbel-type instruments (e.g., Alberini *et al.*, 2019). Of the 1,068 Tier 4 estimates from 91 studies in the full corpus, only 3 studies report any recoverable first-stage instrument-strength diagnostic (a first-stage F, Kleibergen-Paap, or comparable statistic), and none does so in a systematically coded field; we therefore cannot audit instrument strength across the tier and note the coverage rather than adjudicate it. Tier 5: panel fixed effects, cointegration systems, and structural demand systems without price instrumentation (e.g., Bernstein & Griffin, 2006). Tier 6: naive OLS (ordinary least squares) or time-series regressions of quantity on price; Barnes *et al.* (1981), for instance, report such an estimate alongside an instrumented one and trace the gap between them to the simultaneity that block pricing induces between price and quantity. Estimates whose strategy cannot be determined are excluded from ladder analyses (878 estimates spanning 61 studies, of which 91 carry a usable Marshallian-equivalent estimate and so enter the headline pool that feeds every non-ladder analysis: the horizon path, the two-clock test, and the publication-bias battery). The full per-study mapping ships with the replication package. In the short-run headline sample the tiers contain 81 (design-based), 549 (IV), 655 (FE/structural), and 316 (naive) estimates. Coding in this literature, ours included, has historically prioritized each study's preferred estimates;

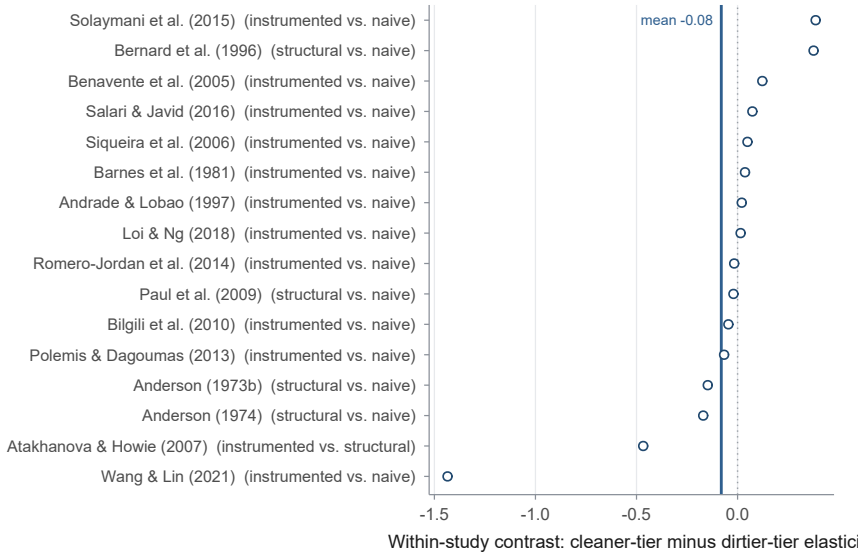
the corpus carries a `preferred` flag (1,204 estimates), and all headline results are robust to restricting the sample to these flagged estimates (Section 4).

The tier assignments were made by three coders from each study's description of its price variation, and the full dataset then underwent a study-by-study audit against the primary sources, including a re-check of the estimates drawn from the earlier compilation of Dahl (2011). Three features make the classification dependable. First, the complete per-study tier mapping ships with the replication package, so every assignment can be checked against the primary source. Second, the classification that carries the argument, design-based (Tiers 1–3) versus not, turns on observable, low-discretion features of a study's design (random assignment, a mandated or plausibly exogenous price change, or an explicit untreated comparison group) rather than on a subjective judgment of estimator quality. Third, the headline analyses collapse the six tiers into four groups (design-based, instrumented, fixed-effects/structural, and naive), so boundary calls within a group do not bear on the design-based-versus-naive contrast that anchors the paper.

Appendix S4 Within-study contrasts and robustness

The imputed data-year offset is immaterial to the calendar-time results. Dropping all 23 estimates with no reported data years (assigned publication year minus three in the baseline) and re-estimating leaves the composition-adjusted short-run data-clock trend at +0.0168 per decade, against the reported +0.017 (unchanged across every horizon and specification at the reported precision); the short- and long-run two-clock data-clock coefficients are likewise unchanged at −0.0334 and −0.0919.

Figure S4: Within-study tier contrasts

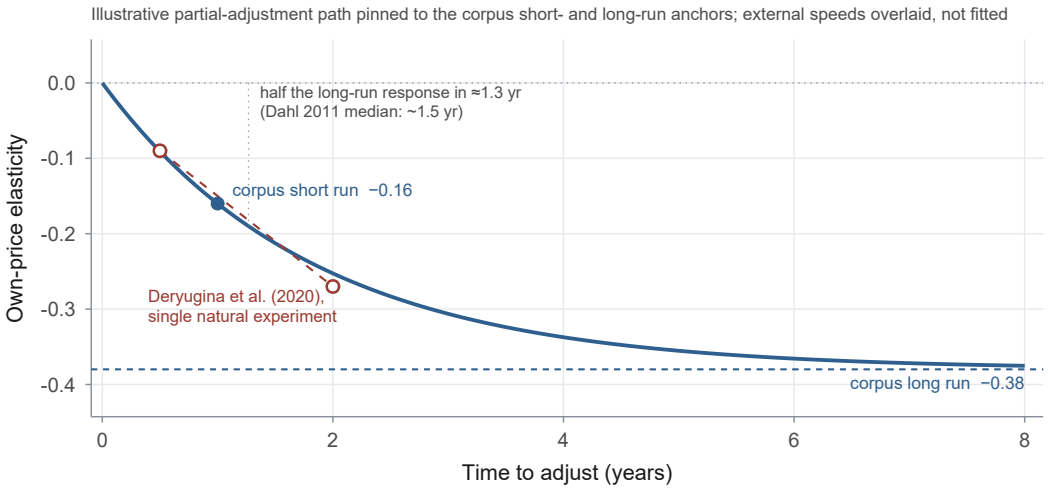


Notes: For each of the sixteen studies reporting estimates on more than one identification tier (tier pair in parentheses), the mean estimate on the cleaner tier minus the mean estimate on the lower tier, same study and data. The mean contrast is −0.081 (cleaner estimates more elastic, driven by IV-vs-naive pairs), with a paired t of −0.79 and positive contrasts in 50% of studies: no significant systematic within-study shift, consistent with the IV overshoot in Table 1.

The adjustment path as a continuous curve. The horizon result of Section 5 is a three-point gradient, short, intermediate, and long run. A single smooth adjustment path reproduces

it (Figure S5). We take a textbook partial-adjustment form $\varepsilon(t) = \varepsilon_{\mathrm{LR}}\,(1 - e^{-\lambda t})$, fix the long-run asymptote at the corrected value (-0.38), and set the speed λ so the curve reaches the corrected short-run value (-0.16) at one year. Nothing in this path is fitted to dynamics in our data, which record only discrete horizon flags, yet it lands on two independent external benchmarks: the median speed of adjustment across dynamic models, half the long-run response in about a year and a half (Dahl, 2011), and the within-experiment trajectory of Deryugina *et al.* (2020), whose six-month and two-year points fall almost on the curve. We read the agreement as corroboration that the horizon gradient is a genuine speed of adjustment rather than composition across horizon bins. We present the curve as an illustration only: it imposes a functional form the rest of the paper avoids, and a true meta-analytic impulse response would require each dynamic study's own adjustment coefficient, which the final dataset does not record.

Figure S5: The adjustment path as a continuous curve (illustrative)



Notes: Illustrative partial-adjustment path $\varepsilon(t) = \varepsilon_{\mathrm{LR}}(1 - e^{-\lambda t})$ with $\varepsilon_{\mathrm{LR}} = -0.38$ (the corrected long-run elasticity) and λ set so $\varepsilon(1\text{ yr}) = -0.16$ (the corrected short-run elasticity); the implied half-life of adjustment is about 1.3 years. Overlaid, not fitted: the median dynamic-model adjustment speed of Dahl (2011) (half the long-run response in about 1.5 years) and the natural-experiment path of Deryugina *et al.* (2020) (-0.09 at six months, -0.27 at two years). The curve illustrates the discrete gradient of Figure 2; it is not a structural estimate.

Sample robustness: the pure-Marshallian core. The baseline pools directly reported Marshallian estimates with Slutsky-converted compensated (Hicksian) ones (Section 2). Restricting instead to the 2,579 directly reported Marshallian estimates alone (from 272 studies) leaves the picture unchanged: short-run PET -0.172 (se 0.028), long-run PET -0.356 (se 0.041), funnel asymmetry $p < 0.001$ in both. The calendar-time trend is unchanged as well: on this native-Marshallian sample the fully adjusted short-run data-clock coefficient is $+0.043$ per decade ($p = 0.234$), the same sign and the same statistical indistinguishability from zero as the pooled headline, and the time-of-use coefficient stays positive and significant ($+0.19$, $p < 0.001$). The harmonization is numerically inconsequential in either direction: because electricity's budget share is small the Slutsky shift is about 0.012 per estimate, varying the budget share s over $[0.02, 0.08]$ moves the short-run corrected value by less than 0.01, and the *converted-from-Hicksian* moderator in the model average (Supplementary appendix S6) is robustly selected only in the long-run model (posterior mean -0.18) and does not overturn the pooled short- and long-run corrections. One-preferred-estimate-per-study and sample-size weighting (no standard errors in the weights) likewise leave the short-run correction in the same range. The long-run

ladder preserves the short-run IV overshoot (design-based cell too thin to report): IV −0.41 against FE −0.33, with naive at −0.23 adjusted.

The estimates held out of the baseline. The 1,396 estimates held out of the Marshallian-equivalent baseline (Supplementary appendix S1 gives their composition) are statistically indistinguishable from the estimation sample: Table S5 shows their raw and bias-corrected means and their funnel asymmetry track the baseline, and in a pooled funnel regression neither the level of the held-out estimates (−0.03, $p = 0.62$) nor their funnel slope (−0.01, $p = 0.97$) differs from it. Re-running the correction on the full corpus of 4,720 leaves the bias-corrected elasticity at −0.25, the value the baseline delivers. The restriction to the estimation sample is a matter of placing estimates on a common scale, not a selection that shapes the result.

Appendix S5 Publication bias: the full battery

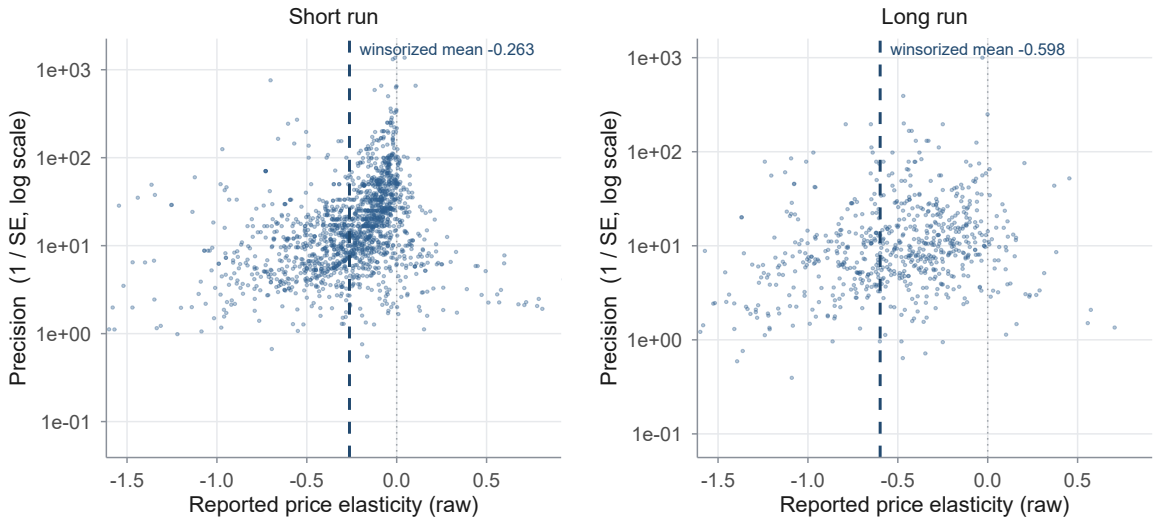
The main text corrects the short-run elasticity with a compact battery (Table 2); here we re-run the *full* battery: the linear FAT-PET family (Panel A), the nonlinear correctors, including the Bom–Rachinger endogenous kink (Panel B), and the endogeneity- and selection-robust estimators (MAIVE, p-uniform*, and a selection model; Panel C), on each sample the paper leans on, so the correction can be read off estimator by estimator. Figure S6 plots the raw short- and long-run estimates against their precision. Table S6 runs the battery on the short-run and long-run headline samples; Table S7 and Table S8 split the short-run sample by identification tier (naive and panel fixed-effects; instrumented and design-based); and Table S9 splits it by price regime (time-of-use tariffs and marginal price). The message is broadly uniform: in nearly every cell the correctors pull the reported mean toward zero, the funnel-asymmetry (FAT) slope is negative wherever the sample is observational, and the least-selected cell, the design-based sample (baseline FAT slope +0.029), carries one of the least price-responsive corrected values; the time-of-use sample is strongly selected (its FAT slope is among the largest in the battery) yet its corrected value is still near zero, which strengthens rather than weakens the low-elasticity conclusion. One further corrector, RTMA (right-truncated meta-analysis; Mathur, 2024), is not identified in this literature and we do not report it: nearly every estimate is *affirmative* (statistically significant with the expected negative sign), so the small residue of non-affirmative estimates on which the right-truncated likelihood relies is too thin to pin down the model, which then over-corrects implausibly past the raw mean (to +0.26 in the short run).

Table S5: Estimates held out of the estimation sample

	N	Studies	Mean	Median	PET (SE)	FAT
Baseline (estimated)	3,324	366	−0.38	−0.28	−0.24 (0.02)	−0.92
Held out of estimation:						
Not Marshallian-convertible	859	77	−0.41	−0.24	−0.27 (0.05)	−0.96
Inverted-formula	420	32	−0.47	−0.31	−0.37 (0.10)	−0.60
Imputed standard error	117	1	−0.11	−0.09	–	–
All held out	1,396	110	−0.40	−0.23	−0.27 (0.05)	−0.93
Full corpus	4,720	462	−0.39	−0.26	−0.25 (0.02)	−0.93

Notes: Own-price elasticities (1%-winsorized), all horizons pooled. "Not Marshallian-convertible" estimates are compensated (Hicksian) elasticities that lack a usable income elasticity for the Slutsky conversion. PET is the FAT-PET bias-corrected mean with study-clustered standard error; FAT is the funnel-asymmetry slope. The imputed-standard-error group is a single study, so its clustered FAT-PET is not identified. In a pooled funnel regression with a held-out indicator, the held-out estimates differ from the baseline neither in level (−0.03, $p = 0.62$) nor in funnel slope (−0.01, $p = 0.97$).

Figure S6: Funnel plots of the short-run and long-run elasticities



Notes: Raw reported own-price elasticities against their precision (1/SE), for the short-run and long-run headline samples; the dashed line marks the 1%-winsorized simple mean of each sample (−0.26 short run, −0.59 long run). The leftward asymmetry, a thinning of imprecise estimates on the positive side, is the funnel signature of selection against small or wrong-signed elasticities. The vertical axis (precision, 1/SE) is drawn on a logarithmic scale, which spreads the dense cluster of imprecise estimates near the bottom and makes the asymmetry easier to see; the horizontal axis is clipped to $[-1.5, 0.8]$ for legibility, and a small number of outliers fall outside the frame but enter every calculation.

Table S6: Publication-bias battery: short-run and long-run elasticities

Part 1: Short-run elasticities					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.851***	-2.055***	-0.953***	-1.281***	-1.171***
(Standard error)	(0.154)	(0.708)	(0.072)	(0.128)	(0.202)
Effect beyond bias	-0.163***	-0.093***	-0.158***	-0.143***	-0.124***
(Constant)	(0.025)	(0.034)	(0.007)	(0.019)	(0.024)
Observations	1,647	1,647	1,647	1,647	1,647
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.104***	-0.111***	-0.231***	-0.239***	-0.113***
	(0.000)	(0.000)	(0.020)	(0.019)	(0.029)
Observations	1,647	1,647	1,647	1,647	1,647
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.145***	-0.181***	-0.155***
			(0.027)	(0.022)	(0.011)
First-stage F (MAIVE)				23	
Observations			1,647	1,647	1,647
Part 2: Long-run elasticities					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.993***	-1.337	-1.055***	-1.095***	-0.979***
(Standard error)	(0.160)	(1.096)	(0.104)	(0.160)	(0.220)
Effect beyond bias	-0.377***	-0.359***	-0.365***	-0.351***	-0.380***
(Constant)	(0.036)	(0.091)	(0.018)	(0.046)	(0.045)
Observations	723	723	723	723	723
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.372***	-0.376***	-0.498***	-0.424***	-0.378***
	(0.001)	(0.001)	(0.032)	(0.114)	(0.076)
Observations	723	723	723	723	723
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.385***		-0.420***
			(0.052)		(0.023)
First-stage F (MAIVE)				< 1	
Observations			723	723	723

Notes: Publication-bias–corrected mean price elasticity from each estimator, on the 1%-winsorized headline sample. *Panel A* is the FAT-PET meta-regression $\varepsilon_{is} = \varepsilon_0 + \gamma\, SE(\varepsilon_{is}) + u_{is}$: "Publication bias" is the slope γ and "Effect beyond bias" the intercept ε_0 (each coefficient's standard error below it); the columns are OLS unweighted, FE inverse-variance weighted, RE random effects (DerSimonian–Laird), Study one median estimate per study, and Precision inverse-SE weighted. Standard errors are two-way clustered by study and by underlying database wherever a regression is estimated (OLS, FE, Precision, PEESE, and EK); RE reports the DerSimonian–Laird standard error and the study-level techniques are already one observation per study. *Panel B*: PEESE (Stanley & Doucouliagos, 2014), WAAP (Ioannidis *et al.*, 2017), the top-10% estimator, STEM (Furukawa, 2019), and EK, the endogenous-kink meta-regression of Bom & Rachinger (2019) whose kink point $a_1 = |\varepsilon_0|/1.96$ is iterated to convergence. *Panel C*: MAIVE (Irsova *et al.*, 2025) instruments the estimate variance with the inverse sample size (top 1% of SEs dropped) and is reported only where the first-stage instrument is adequate ($F \geq 10$; otherwise only the F is shown); p-uniform* (van Aert & van Assen, 2026) works from study medians. "Selection" is a step-function selection model (a Vevea–Hedges / Andrews–Kasy-type maximum-likelihood correction (Vevea & Hedges, 1995; Andrews & Kasy, 2019) with a cut at the one-sided 2.5% threshold). "Observations" is the number of estimates in the sample, all of which enter every test; the study-level techniques (Study, STEM, p-uniform*) then aggregate them to one value per study, and WAAP and the top-10% estimator average a precision-selected subset of the same estimates. ***, **, and * denote statistical significance at the 1%, 5%, and 10% level.

Table S7: Publication-bias battery: naive and panel fixed-effects tiers

Part 1: Naive					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.750***	-3.187*	-0.736***	-1.227***	-1.330***
(Standard error)	(0.204)	(1.752)	(0.170)	(0.206)	(0.418)
Effect beyond bias	-0.245***	-0.081	-0.248***	-0.188***	-0.152**
(Constant)	(0.049)	(0.093)	(0.023)	(0.044)	(0.073)
Observations	316	316	316	316	316
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.083***	-0.106***	-0.318***	-0.300***	-0.106
	(0.001)	(0.001)	(0.041)	(0.052)	(0.085)
Observations	316	316	316	316	316
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.199***		-0.226***
			(0.066)		(0.033)
First-stage F (MAIVE)				< 1	
Observations			316	316	316
Part 2: Panel FE / structural					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.965***	0.032	-0.885***	-1.589***	-0.683**
(Standard error)	(0.174)	(1.165)	(0.124)	(0.146)	(0.316)
Effect beyond bias	-0.088***	-0.150**	-0.098***	-0.085***	-0.123***
(Constant)	(0.026)	(0.069)	(0.014)	(0.023)	(0.041)
Observations	655	655	655	655	655
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.157***	-0.151***	-0.164***	-0.028	-0.149***
	(0.001)	(0.001)	(0.020)	(0.134)	(0.057)
Observations	655	655	655	655	655
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.128***		-0.115***
			(0.031)		(0.013)
First-stage F (MAIVE)				1	
Observations			655	655	655

Notes: Publication-bias–corrected mean price elasticity from each estimator, on the 1%-winsorized headline sample. *Panel A* is the FAT-PET meta-regression $\varepsilon_{is} = \varepsilon_0 + \gamma\, SE(\varepsilon_{is}) + u_{is}$: "Publication bias" is the slope γ and "Effect beyond bias" the intercept ε_0 (each coefficient's standard error below it); the columns are OLS unweighted, FE inverse-variance weighted, RE random effects (DerSimonian–Laird), Study one median estimate per study, and Precision inverse-SE weighted. Standard errors are two-way clustered by study and by underlying database wherever a regression is estimated (OLS, FE, Precision, PEESE, and EK); RE reports the DerSimonian–Laird standard error and the study-level techniques are already one observation per study. *Panel B*: PEESE (Stanley & Doucouliagos, 2014), WAAP (Ioannidis *et al.*, 2017), the top-10% estimator, STEM (Furukawa, 2019), and EK, the endogenous-kink meta-regression of Bom & Rachinger (2019) whose kink point $a_1 = |\varepsilon_0|/1.96$ is iterated to convergence. *Panel C*: MAIVE (Irsova *et al.*, 2025) instruments the estimate variance with the inverse sample size (top 1% of SEs dropped) and is reported only where the first-stage instrument is adequate ($F \geq 10$; otherwise only the F is shown); p-uniform* (van Aert & van Assen, 2026) works from study medians. "Selection" is a step-function selection model (a Vevea–Hedges / Andrews–Kasy-type maximum-likelihood correction (Vevea & Hedges, 1995; Andrews & Kasy, 2019) with a cut at the one-sided 2.5% threshold). "Observations" is the number of estimates in the sample, all of which enter every test; the study-level techniques (Study, STEM, p-uniform*) then aggregate them to one value per study, and WAAP and the top-10% estimator average a precision-selected subset of the same estimates. ***, **, and * denote statistical significance at the 1%, 5%, and 10% level.

Table S8: Publication-bias battery: instrumented and design-based tiers

Part 1: Instrumented (IV/GMM)					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-1.037***	-3.819***	-1.462***	-1.722***	-1.841***
(Standard error)	(0.241)	(0.695)	(0.109)	(0.361)	(0.375)
Effect beyond bias	-0.203***	-0.063***	-0.177***	-0.119***	-0.124***
(Constant)	(0.039)	(0.023)	(0.009)	(0.039)	(0.036)
Observations	549	549	549	549	549
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.074***	-0.102***	-0.276***	-0.079***	-0.102***
	(0.001)	(0.001)	(0.030)	(0.022)	(0.028)
Observations	549	549	549	549	549
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.118**	-0.278***	-0.164***
			(0.056)	(0.042)	(0.020)
First-stage F (MAIVE)				17	
Observations			549	549	549
Part 2: Design-based (RCT/nat. exp./DID)					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	0.029	0.467	-0.486**	-0.556	-0.185
(Standard error)	(0.596)	(1.055)	(0.204)	(0.572)	(0.612)
Effect beyond bias	-0.095*	-0.084**	-0.046***	-0.117	-0.066**
(Constant)	(0.050)	(0.036)	(0.011)	(0.103)	(0.030)
Observations	81	81	81	81	81
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.087***	-0.080***	-0.114**	-0.108***	-0.077***
	(0.003)	(0.002)	(0.056)	(0.038)	(0.028)
Observations	81	81	81	81	81
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.116	-0.115***	-0.066***
			(0.078)	(0.022)	(0.014)
First-stage F (MAIVE)				123	
Observations			81	81	81

Notes: Publication-bias–corrected mean price elasticity from each estimator, on the 1%-winsorized headline sample. *Panel A* is the FAT-PET meta-regression $\varepsilon_{is} = \varepsilon_0 + \gamma\, SE(\varepsilon_{is}) + u_{is}$: "Publication bias" is the slope γ and "Effect beyond bias" the intercept ε_0 (each coefficient's standard error below it); the columns are OLS unweighted, FE inverse-variance weighted, RE random effects (DerSimonian–Laird), Study one median estimate per study, and Precision inverse-SE weighted. Standard errors are two-way clustered by study and by underlying database wherever a regression is estimated (OLS, FE, Precision, PEESE, and EK); RE reports the DerSimonian–Laird standard error and the study-level techniques are already one observation per study. *Panel B*: PEESE (Stanley & Doucouliagos, 2014), WAAP (Ioannidis *et al.*, 2017), the top-10% estimator, STEM (Furukawa, 2019), and EK, the endogenous-kink meta-regression of Bom & Rachinger (2019) whose kink point $a_1 = |\varepsilon_0|/1.96$ is iterated to convergence. *Panel C*: MAIVE (Irsova *et al.*, 2025) instruments the estimate variance with the inverse sample size (top 1% of SEs dropped) and is reported only where the first-stage instrument is adequate ($F \geq 10$; otherwise only the F is shown); p-uniform* (van Aert & van Assen, 2026) works from study medians. "Selection" is a step-function selection model (a Vevea–Hedges / Andrews–Kasy-type maximum-likelihood correction (Vevea & Hedges, 1995; Andrews & Kasy, 2019) with a cut at the one-sided 2.5% threshold). "Observations" is the number of estimates in the sample, all of which enter every test; the study-level techniques (Study, STEM, p-uniform*) then aggregate them to one value per study, and WAAP and the top-10% estimator average a precision-selected subset of the same estimates. ***, **, and * denote statistical significance at the 1%, 5%, and 10% level.

Table S9: Publication-bias battery: time-of-use and marginal-price subsamples

Part 1: Time-of-use tariff					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.517**	-1.754***	-1.241***	-0.907*	-1.001***
(Standard error)	(0.262)	(0.490)	(0.120)	(0.484)	(0.372)
Effect beyond bias	-0.091**	0.001	-0.022***	-0.147*	-0.018
(Constant)	(0.037)	(0.012)	(0.008)	(0.083)	(0.015)
Observations	210	210	210	210	210
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.005***	-0.011***	-0.154***	-0.051	0.001
	(0.001)	(0.001)	(0.056)	(0.038)	(0.012)
Observations	210	210	210	210	210
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.228***	-0.090***	-0.075***
			(0.076)	(0.032)	(0.020)
First-stage F (MAIVE)				15	
Observations			210	210	210
Part 2: Marginal price					
Panel A: Linear techniques					
	OLS	FE	RE	Study	Precision
Publication bias	-0.969***	-2.945***	-1.530***	-1.239***	-1.396***
(Standard error)	(0.297)	(0.847)	(0.098)	(0.193)	(0.349)
Effect beyond bias	-0.148***	-0.027	-0.114***	-0.135***	-0.080**
(Constant)	(0.057)	(0.022)	(0.008)	(0.047)	(0.031)
Observations	401	401	401	401	401
Panel B: Nonlinear techniques					
	Top 10%	WAAP	PEESE	STEM	EK
Effect beyond bias	-0.041***	-0.049***	-0.230***	-0.076*	-0.049**
	(0.001)	(0.001)	(0.047)	(0.039)	(0.025)
Observations	401	401	401	401	401
Panel C: Endogeneity- and selection-robust techniques					
			p-uniform*	MAIVE	Selection
Effect beyond bias			-0.124**		-0.155***
			(0.061)		(0.022)
First-stage F (MAIVE)				7	
Observations			401	401	401

Notes: Publication-bias–corrected mean price elasticity from each estimator, on the 1%-winsorized headline sample. *Panel A* is the FAT-PET meta-regression $\varepsilon_{is} = \varepsilon_0 + \gamma\, SE(\varepsilon_{is}) + u_{is}$: "Publication bias" is the slope γ and "Effect beyond bias" the intercept ε_0 (each coefficient's standard error below it); the columns are OLS unweighted, FE inverse-variance weighted, RE random effects (DerSimonian–Laird), Study one median estimate per study, and Precision inverse-SE weighted. Standard errors are two-way clustered by study and by underlying database wherever a regression is estimated (OLS, FE, Precision, PEESE, and EK); RE reports the DerSimonian–Laird standard error and the study-level techniques are already one observation per study. *Panel B*: PEESE (Stanley & Doucouliagos, 2014), WAAP (Ioannidis *et al.*, 2017), the top-10% estimator, STEM (Furukawa, 2019), and EK, the endogenous-kink meta-regression of Bom & Rachinger (2019) whose kink point $a_1 = |\varepsilon_0|/1.96$ is iterated to convergence. *Panel C*: MAIVE (Irsova *et al.*, 2025) instruments the estimate variance with the inverse sample size (top 1% of SEs dropped) and is reported only where the first-stage instrument is adequate ($F \geq 10$; otherwise only the F is shown); p-uniform* (van Aert & van Assen, 2026) works from study medians. "Selection" is a step-function selection model (a Vevea–Hedges / Andrews–Kasy-type maximum-likelihood correction (Vevea & Hedges, 1995; Andrews & Kasy, 2019) with a cut at the one-sided 2.5% threshold). "Observations" is the number of estimates in the sample, all of which enter every test; the study-level techniques (Study, STEM, p-uniform*) then aggregate them to one value per study, and WAAP and the top-10% estimator average a precision-selected subset of the same estimates. ***, **, and * denote statistical significance at the 1%, 5%, and 10% level.

Appendix S6 Heterogeneity: Bayesian model averaging

The identification ladder and the two-clock test isolate the two dimensions the paper cares about most; the literature varies along many others, and this appendix maps that variation with Bayesian model averaging (BMA), computed with the `BMS` package (Zeugner & Feldkircher, 2015).[1] The exercise is descriptive and deliberately secondary to the trend and ladder findings: it asks which of the coded study characteristics move the reported elasticity once all of them compete for inclusion, and it doubles as a robustness check on the publication-bias term, which here must survive against every moderator at once. Because the paper's spine is the adjustment horizon, we run the averaging on three samples (the full headline sample and, separately, its short-run and long-run subsamples) with variable definitions and summary statistics in Table S10 (full sample), Table S11 (short run) and Table S12 (long run). Table S13 reports the integrated averaging beside a two-way (study × database) clustered OLS check and a frequentist model average; Table S14 and Table S15 report the horizon-specific models; Table S16 the sensitivity to the priors and Table S17 the convergence diagnostics; Figure S7 shows the model-inclusion picture for the three samples and Figure S8 the regressor correlation matrix.

The response variable is the 1%-winsorized price elasticity and the publication-bias term is its 1%-winsorized standard error, as throughout the paper. The moderators, their reference categories, and the log transforms follow the variable list that ships with the final dataset. Because the identification, model-form, data-structure and price-measurement indicators are strongly collinear (a naive static reduced-form model on yearly average prices is one bundle, a dynamic instrumented panel another), we average under the *dilution* prior of George (2010), which down-weights models whose regressor correlation matrix is near-singular and so prices in the redundancy the single-regression approach ignores; the scale-laden moderators (study size, data span, publication year, impact factor, citations) enter in logs. Two-way clustering by study and by underlying database sharpens the frequentist check: many estimates share a data source across studies, and clustering on the database as well as the study nets out that second layer of dependence. The two-way correction changes little, because shared datasets are rare in this corpus: the 366 headline studies draw on 344 distinct underlying data sources, only 17 databases are shared by two or more studies and none by more than four, so study-level clustering already absorbs almost all of the cross-estimate dependence. Re-running the headline inferences with two-way (study × database) clustering leaves them materially unchanged: the short-run PET is -0.16 with study-clustered and two-way-clustered standard errors identical to three decimals (0.025), the design-versus-naive ladder contrast moves from $p = 0.006$ to $p = 0.007$, and the composition-adjusted data-clock trend stays null ($p = 0.53$). The lone coefficient in which the second clustering dimension is visible is the pooled design-based term of the integrated model (Table S13), borderline at $p = 0.08$, a reflection of its larger moderator set and pooled horizons rather than of the clustering. Nine country-level covariate values (daylight and population) are median-imputed; the elasticity and its standard error are never imputed.

[1] The replication package (data and code) for the full paper, including all appendices, is available at `https://meta-analysis.cz/electricity`.

Table S10: Definition and summary statistics of the model-averaging variables: full sample

Variable	Description	Mean	SD
Price elasticity	The own-price elasticity of electricity demand (1%-winsorized), the response variable.	-0.39	0.44
Standard error	The standard error of the elasticity estimate (1%-winsorized); its coefficient measures publication selection.	0.15	0.20
Horizon			
Short run	=1 if the estimate is a short-run (within about one year) elasticity.	0.50	0.50
Long run	=1 if the estimate is a long-run (beyond about five years; capital fully adjusted) elasticity.	0.22	0.41
Intermediate run	=1 if the estimate is an intermediate-run (about 2–5 years) elasticity (reference category).	0.29	0.45
Estimate type			
Converted from Hicksian	=1 if the estimate is a compensated (Hicksian) own-price elasticity converted to a Marshallian footing via the Slutsky relation $\varepsilon_M = \varepsilon_H - s\eta$; =0 for a directly reported Marshallian estimate. Controls for any residual difference between the converted and the directly reported evidence.	0.22	0.42
Identification			
Design-based	=1 if a design-based strategy identifies the price response: a randomized or mandated pricing experiment, a natural experiment, or a difference-in-differences design (tiers 1–3).	0.05	0.21
Instrumented	=1 if the price is instrumented (IV, two- or three-stage least squares, or GMM; tier 4).	0.27	0.45
Panel FE / structural	=1 if identification rests on panel fixed effects or a structural demand system, without instrumentation (tier 5).	0.43	0.50
Naive	=1 if a naive regression with no price-identification strategy is used (tier 6; the reference category, which also absorbs estimates whose strategy could not be classified).	0.22	0.41
Estimate aggregation			
Aggregation: country	=1 if the data are aggregated to the country (or larger) level.	0.26	0.44
Aggregation: micro	=1 if the data are micro (household, plant or firm-level).	0.27	0.44
Aggregation: region	=1 if the data are aggregated to a sub-country region or city (reference category).	0.30	0.46
Sector			
Residential demand	=1 if the estimate refers to residential electricity demand.	0.62	0.49
Industrial demand	=1 if the estimate refers to industrial electricity demand.	0.19	0.39
Commercial demand	=1 if the estimate refers to commercial or mixed demand (reference category).	0.09	0.28
Data characteristics			
Time-series data	=1 if the study uses time-series data.	0.34	0.47
Cross-section data	=1 if the study uses cross-sectional data.	0.12	0.33
Panel data	=1 if the study uses panel data (reference category).	0.54	0.50
Yearly data	=1 if the data are of yearly frequency.	0.66	0.47
Sub-yearly data	=1 if the data are of monthly or quarterly frequency (reference category).	0.28	0.45
Data span	The logarithm of one plus the number of years the dataset covers.	2.35	1.00
Study size	The logarithm of the number of observations used in the estimation.	6.13	2.54
Price and tariff			
Average price	=1 if the price is measured as average price (revenue over quantity).	0.64	0.48

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Table S10 continued

Variable	Description	Mean	SD
Marginal price	=1 if the price is measured as marginal price.	0.24	0.43
Other price	=1 if another price concept is used (lump-sum, Shin, undefined; reference category).	0.08	0.27
Increasing-block tariff	=1 if an increasing-block tariff is in force.	0.19	0.39
Decreasing-block tariff	=1 if a decreasing-block tariff is in force.	0.12	0.32
Time-of-use tariff	=1 if a time-of-use rate (different unit prices by time block) is used.	0.14	0.34
Flat tariff	=1 if a flat tariff is in force (reference category).	0.04	0.20
Demand controls			
Control: demographics	=1 if the demand equation controls for demographic variation.	0.36	0.48
Control: temperature	=1 if the demand equation controls for temperature or degree-days.	0.50	0.50
Control: appliance stock	=1 if the demand equation controls for the stock of electrical appliances.	0.22	0.41
Control: other fuels	=1 if the demand equation controls for prices of substitute fuels.	0.41	0.49
Model specification			
Reduced form	=1 if a single-equation reduced-form model is estimated.	0.48	0.50
Structural form	=1 if a structural (multi-equation) model is estimated (reference category).	0.50	0.50
Static model	=1 if the model is static (no lagged demand).	0.32	0.47
Dynamic model	=1 if the model is dynamic, with lagged adjustment (reference category).	0.68	0.47
ARDL	=1 if an autoregressive distributed-lag (partial-adjustment) model is used.	0.04	0.20
Lagged endogenous	=1 if a lagged-endogenous model is used.	0.19	0.40
Linear demand	=1 if the demand function is linear.	0.13	0.34
Double-log demand	=1 if the demand function is double-logarithmic.	0.80	0.40
Semi-log demand	=1 if the demand function is semi-logarithmic or Box-Cox (reference category).	0.01	0.11
Geography and context			
United States	=1 if the estimate is for the United States.	0.41	0.49
Europe	=1 if the estimate is for a European country.	0.18	0.38
Other country	=1 if the estimate is for a country outside the US and Europe (reference category).	0.42	0.49
Daylight hours	The average length of the longest day for the country under examination (source: Moesen, 2010).	14.76	1.58
Population	The logarithm of the population of the country or entity (source: World Bank, 2025).	18.36	1.58
Publication characteristics			
Publication year	The logarithm of one plus the number of years since the earliest study in the sample.	3.91	0.35
Impact factor	The logarithm of one plus the journal's RePEc recursive impact factor; 0 for unpublished work.	0.26	0.55
Citations	The logarithm of one plus the study's mean annual Google Scholar citation count.	0.87	0.67

Notes: Mean and standard deviation over all 3,324 headline estimates (Marshallian-equivalent: directly reported or Slutsky-converted, usable standard error) from 366 studies. Binary indicators are reported as means (shares); the elasticity and its standard error are the 1%-winsorized values used throughout. Reference categories are shown for completeness but drop from the averaging. The identification reference category (Naive) totals 821 estimates, of which 730 are labeled naive (tier 6) and 91 could not be classified; the averaging does not distinguish between the two. Scale-laden variables enter the averaging in logs, as noted in their descriptions. Country-level covariates are drawn from external sources: GDP per capita (current US$) and population from the World Bank's *World Development Indicators* (World Bank, 2025); surface temperature (degrees Celsius) from the Berkeley Earth record (Berkeley Earth, 2015); and the length of the longest day from the daylight calculator of Moesen (2010). Study-level bibliometrics are citation counts from Google Scholar and journal impact factors from the RePEc recursive ranking. The replication package's codebook maps every coded variable to its source.

Table S11: Definition and summary statistics of the model-averaging variables: short-run estimates

Variable	Description	Mean	SD
Price elasticity	The own-price elasticity of electricity demand (1%-winsorized), the response variable.	-0.27	0.33
Standard error	The standard error of the elasticity estimate (1%-winsorized); its coefficient measures publication selection.	0.12	0.17
Horizon			
Short run	=1 if the estimate is a short-run (within about one year) elasticity.	1.00	0.00
Long run	=1 if the estimate is a long-run (beyond about five years; capital fully adjusted) elasticity.	0.00	0.00
Intermediate run	=1 if the estimate is an intermediate-run (about 2–5 years) elasticity (reference category).	0.00	0.00
Estimate type			
Converted from Hicksian	=1 if the estimate is a compensated (Hicksian) own-price elasticity converted to a Marshallian footing via the Slutsky relation $\varepsilon_M = \varepsilon_H - s\eta$; =0 for a directly reported Marshallian estimate. Controls for any residual difference between the converted and the directly reported evidence.	0.18	0.39
Identification			
Design-based	=1 if a design-based strategy identifies the price response: a randomized or mandated pricing experiment, a natural experiment, or a difference-in-differences design (tiers 1–3).	0.05	0.22
Instrumented	=1 if the price is instrumented (IV, two- or three-stage least squares, or GMM; tier 4).	0.33	0.47
Panel FE / structural	=1 if identification rests on panel fixed effects or a structural demand system, without instrumentation (tier 5).	0.40	0.49
Naive	=1 if a naive regression with no price-identification strategy is used (tier 6; the reference category, which also absorbs estimates whose strategy could not be classified).	0.19	0.39
Estimate aggregation			
Aggregation: country	=1 if the data are aggregated to the country (or larger) level.	0.24	0.43
Aggregation: micro	=1 if the data are micro (household, plant or firm-level).	0.26	0.44
Aggregation: region	=1 if the data are aggregated to a sub-country region or city (reference category).	0.34	0.47
Sector			
Residential demand	=1 if the estimate refers to residential electricity demand.	0.64	0.48
Industrial demand	=1 if the estimate refers to industrial electricity demand.	0.13	0.34
Commercial demand	=1 if the estimate refers to commercial or mixed demand (reference category).	0.10	0.30
Data characteristics			
Time-series data	=1 if the study uses time-series data.	0.34	0.47
Cross-section data	=1 if the study uses cross-sectional data.	0.08	0.27
Panel data	=1 if the study uses panel data (reference category).	0.59	0.49
Yearly data	=1 if the data are of yearly frequency.	0.65	0.48
Sub-yearly data	=1 if the data are of monthly or quarterly frequency (reference category).	0.27	0.44
Data span	The logarithm of one plus the number of years the dataset covers.	2.40	0.98
Study size	The logarithm of the number of observations used in the estimation.	6.55	2.56

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Table S11 continued

Variable	Description	Mean	SD
Price and tariff			
Average price	=1 if the price is measured as average price (revenue over quantity).	0.63	0.48
Marginal price	=1 if the price is measured as marginal price.	0.24	0.43
Other price	=1 if another price concept is used (lump-sum, Shin, undefined; reference category).	0.07	0.25
Increasing-block tariff	=1 if an increasing-block tariff is in force.	0.25	0.43
Decreasing-block tariff	=1 if a decreasing-block tariff is in force.	0.07	0.26
Time-of-use tariff	=1 if a time-of-use rate (different unit prices by time block) is used.	0.13	0.33
Flat tariff	=1 if a flat tariff is in force (reference category).	0.04	0.20
Demand controls			
Control: demographics	=1 if the demand equation controls for demographic variation.	0.39	0.49
Control: temperature	=1 if the demand equation controls for temperature or degree-days.	0.58	0.49
Control: appliance stock	=1 if the demand equation controls for the stock of electrical appliances.	0.21	0.40
Control: other fuels	=1 if the demand equation controls for prices of substitute fuels.	0.42	0.49
Model specification			
Reduced form	=1 if a single-equation reduced-form model is estimated.	0.55	0.50
Structural form	=1 if a structural (multi-equation) model is estimated (reference category).	0.43	0.50
Static model	=1 if the model is static (no lagged demand).	0.21	0.40
Dynamic model	=1 if the model is dynamic, with lagged adjustment (reference category).	0.79	0.41
ARDL	=1 if an autoregressive distributed-lag (partial-adjustment) model is used.	0.03	0.18
Lagged endogenous	=1 if a lagged-endogenous model is used.	0.32	0.47
Linear demand	=1 if the demand function is linear.	0.09	0.29
Double-log demand	=1 if the demand function is double-logarithmic.	0.88	0.33
Semi-log demand	=1 if the demand function is semi-logarithmic or Box-Cox (reference category).	0.01	0.10
Geography and context			
United States	=1 if the estimate is for the United States.	0.38	0.49
Europe	=1 if the estimate is for a European country.	0.17	0.37
Other country	=1 if the estimate is for a country outside the US and Europe (reference category).	0.47	0.50
Daylight hours	The average length of the longest day for the country under examination (source: Moesen, 2010).	14.66	1.64
Population	The logarithm of the population of the country or entity (source: World Bank, 2025).	18.32	1.56
Publication characteristics			
Publication year	The logarithm of one plus the number of years since the earliest study in the sample.	3.99	0.29
Impact factor	The logarithm of one plus the journal's RePEc recursive impact factor; 0 for unpublished work.	0.26	0.56
Citations	The logarithm of one plus the study's mean annual Google Scholar citation count.	0.91	0.59

Notes: Mean and standard deviation over the 1,647 short-run headline estimates. The horizon indicators are constant on this subsample and drop from the short-run averaging. Binary indicators are reported as means (shares); the elasticity and its standard error are the 1%-winsorized values.

Table S12: Definition and summary statistics of the model-averaging variables: long-run estimates

Variable	Description	Mean	SD
Price elasticity	The own-price elasticity of electricity demand (1%-winsorized), the response variable.	-0.58	0.52
Standard error	The standard error of the elasticity estimate (1%-winsorized); its coefficient measures publication selection.	0.21	0.26
Horizon			
Short run	=1 if the estimate is a short-run (within about one year) elasticity.	0.00	0.00
Long run	=1 if the estimate is a long-run (beyond about five years; capital fully adjusted) elasticity.	1.00	0.00
Intermediate run	=1 if the estimate is an intermediate-run (about 2–5 years) elasticity (reference category).	0.00	0.00
Estimate type			
Converted from Hicksian	=1 if the estimate is a compensated (Hicksian) own-price elasticity converted to a Marshallian footing via the Slutsky relation $\varepsilon_M = \varepsilon_H - s\eta$; =0 for a directly reported Marshallian estimate. Controls for any residual difference between the converted and the directly reported evidence.	0.36	0.48
Identification			
Design-based	=1 if a design-based strategy identifies the price response: a randomized or mandated pricing experiment, a natural experiment, or a difference-in-differences design (tiers 1–3).	0.01	0.07
Instrumented	=1 if the price is instrumented (IV, two- or three-stage least squares, or GMM; tier 4).	0.26	0.44
Panel FE / structural	=1 if identification rests on panel fixed effects or a structural demand system, without instrumentation (tier 5).	0.55	0.50
Naive	=1 if a naive regression with no price-identification strategy is used (tier 6; the reference category, which also absorbs estimates whose strategy could not be classified).	0.17	0.37
Estimate aggregation			
Aggregation: country	=1 if the data are aggregated to the country (or larger) level.	0.37	0.48
Aggregation: micro	=1 if the data are micro (household, plant or firm-level).	0.18	0.38
Aggregation: region	=1 if the data are aggregated to a sub-country region or city (reference category).	0.38	0.49
Sector			
Residential demand	=1 if the estimate refers to residential electricity demand.	0.64	0.48
Industrial demand	=1 if the estimate refers to industrial electricity demand.	0.16	0.36
Commercial demand	=1 if the estimate refers to commercial or mixed demand (reference category).	0.10	0.30
Data characteristics			
Time-series data	=1 if the study uses time-series data.	0.36	0.48
Cross-section data	=1 if the study uses cross-sectional data.	0.17	0.37
Panel data	=1 if the study uses panel data (reference category).	0.47	0.50
Yearly data	=1 if the data are of yearly frequency.	0.82	0.39
Sub-yearly data	=1 if the data are of monthly or quarterly frequency (reference category).	0.12	0.32
Data span	The logarithm of one plus the number of years the dataset covers.	2.74	0.95
Study size	The logarithm of the number of observations used in the estimation.	5.62	2.10

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Table S12 continued

Variable	Description	Mean	SD
Price and tariff			
Average price	=1 if the price is measured as average price (revenue over quantity).	0.81	0.39
Marginal price	=1 if the price is measured as marginal price.	0.07	0.26
Other price	=1 if another price concept is used (lump-sum, Shin, undefined; reference category).	0.10	0.30
Increasing-block tariff	=1 if an increasing-block tariff is in force.	0.21	0.41
Decreasing-block tariff	=1 if a decreasing-block tariff is in force.	0.06	0.24
Time-of-use tariff	=1 if a time-of-use rate (different unit prices by time block) is used.	0.04	0.20
Flat tariff	=1 if a flat tariff is in force (reference category).	0.04	0.20
Demand controls			
Control: demographics	=1 if the demand equation controls for demographic variation.	0.35	0.48
Control: temperature	=1 if the demand equation controls for temperature or degree-days.	0.44	0.50
Control: appliance stock	=1 if the demand equation controls for the stock of electrical appliances.	0.19	0.39
Control: other fuels	=1 if the demand equation controls for prices of substitute fuels.	0.32	0.47
Model specification			
Reduced form	=1 if a single-equation reduced-form model is estimated.	0.40	0.49
Structural form	=1 if a structural (multi-equation) model is estimated (reference category).	0.57	0.50
Static model	=1 if the model is static (no lagged demand).	0.25	0.43
Dynamic model	=1 if the model is dynamic, with lagged adjustment (reference category).	0.74	0.44
ARDL	=1 if an autoregressive distributed-lag (partial-adjustment) model is used.	0.12	0.32
Lagged endogenous	=1 if a lagged-endogenous model is used.	0.05	0.21
Linear demand	=1 if the demand function is linear.	0.05	0.21
Double-log demand	=1 if the demand function is double-logarithmic.	0.93	0.26
Semi-log demand	=1 if the demand function is semi-logarithmic or Box-Cox (reference category).	0.01	0.11
Geography and context			
United States	=1 if the estimate is for the United States.	0.24	0.43
Europe	=1 if the estimate is for a European country.	0.28	0.45
Other country	=1 if the estimate is for a country outside the US and Europe (reference category).	0.49	0.50
Daylight hours	The average length of the longest day for the country under examination (source: Moesen, 2010).	14.64	1.88
Population	The logarithm of the population of the country or entity (source: World Bank, 2025).	18.21	1.79
Publication characteristics			
Publication year	The logarithm of one plus the number of years since the earliest study in the sample.	4.09	0.22
Impact factor	The logarithm of one plus the journal's RePEc recursive impact factor; 0 for unpublished work.	0.40	0.78
Citations	The logarithm of one plus the study's mean annual Google Scholar citation count.	1.08	0.99

Notes: Mean and standard deviation over the 723 long-run headline estimates. The horizon indicators are constant on this subsample and drop from the long-run averaging. Binary indicators are reported as means (shares); the elasticity and its standard error are the 1%-winsorized values.

Table S13: Bayesian model averaging of the price elasticity: integrated headline sample

	Bayesian model averaging (UIP g-prior, dilution)			OLS, two-way clustered (study × database)			FMA (smoothed AIC)	
	P. mean	P. SD	PIP	Coef.	SE	p	Coef.	Wt.
Constant	-0.480	–	1.000	-0.937	0.257	0.000	-0.668	1.000
Standard error	-0.830	0.034	1.000	-0.830	0.100	0.000	-0.834	1.000
Horizon								
Short run	0.114	0.017	1.000	0.096	0.034	0.005	0.114	1.000
Long run	-0.103	0.021	1.000	-0.122	0.039	0.002	-0.101	1.000
Estimate type								
Converted from Hicksian	-0.000	0.003	0.014	-0.011	0.028	0.685	-0.000	0.012
Identification								
Design-based	0.159	0.039	0.994	0.122	0.069	0.080	0.144	0.999
Instrumented	-0.001	0.006	0.019	-0.023	0.038	0.549	-0.000	0.003
Panel FE / structural	0.075	0.019	0.989	0.058	0.030	0.055	0.072	1.000
Estimate aggregation								
Aggregation: country	0.000	0.002	0.010	-0.009	0.028	0.739	0.000	0.002
Aggregation: micro	-0.022	0.031	0.376	-0.039	0.040	0.340	-0.020	0.394
Sector								
Residential demand	-0.012	0.020	0.293	-0.028	0.028	0.310	-0.018	0.460
Industrial demand	0.000	0.004	0.018	0.002	0.036	0.961	0.000	0.014
Data characteristics								
Time-series data	0.003	0.010	0.074	0.022	0.032	0.492	0.004	0.135
Cross-section data	-0.166	0.029	1.000	-0.146	0.046	0.002	-0.154	1.000
Yearly data	-0.125	0.019	1.000	-0.108	0.033	0.001	-0.125	1.000
Data span	0.005	0.012	0.200	0.019	0.021	0.362	0.019	0.664
Study size	-0.017	0.004	0.995	-0.018	0.006	0.005	-0.018	1.000
Price and tariff								
Average price	-0.000	0.002	0.013	0.010	0.034	0.780	-0.000	0.007
Marginal price	0.005	0.016	0.118	0.087	0.055	0.116	0.050	0.874
Increasing-block tariff	0.000	0.002	0.009	0.025	0.035	0.483	0.000	0.002
Decreasing-block tariff	-0.174	0.024	1.000	-0.142	0.054	0.008	-0.178	1.000
Time-of-use tariff	0.000	0.004	0.014	0.038	0.046	0.405	0.001	0.032
Demand controls								
Control: demographics	-0.067	0.016	0.998	-0.057	0.027	0.038	-0.058	1.000
Control: temperature	-0.000	0.002	0.010	-0.012	0.025	0.622	-0.000	0.002
Control: appliance stock	-0.005	0.015	0.118	-0.032	0.034	0.341	-0.025	0.546
Control: other fuels	-0.074	0.016	1.000	-0.062	0.028	0.029	-0.070	1.000
Model specification								
Reduced form	0.014	0.021	0.338	0.041	0.029	0.157	0.024	0.661
Static model	-0.000	0.004	0.016	-0.041	0.035	0.245	-0.002	0.061
ARDL	0.002	0.011	0.034	0.050	0.043	0.251	0.002	0.034
Lagged endogenous	0.101	0.020	1.000	0.083	0.036	0.020	0.095	1.000
Linear demand	0.010	0.023	0.188	0.049	0.052	0.351	0.041	0.701
Double-log demand	-0.004	0.014	0.094	-0.005	0.041	0.904	-0.003	0.067
Geography and context								
United States	0.037	0.030	0.665	0.054	0.034	0.110	0.027	0.512
Europe	0.002	0.010	0.046	0.043	0.041	0.297	0.009	0.184
Daylight hours	0.019	0.005	0.992	0.018	0.007	0.009	0.019	0.999
Population	0.004	0.007	0.313	0.014	0.010	0.150	0.011	0.740
Publication characteristics								
Publication year	0.001	0.007	0.022	0.065	0.050	0.194	0.006	0.088
Impact factor	0.034	0.027	0.665	0.031	0.024	0.190	0.028	0.653
Citations	0.018	0.021	0.474	0.019	0.021	0.352	0.020	0.611
Observations	3,324			3,324			3,324	
Studies / databases	366 / 344							

Notes: The response variable is the 1%-winsorized own-price elasticity of electricity demand. *Bayesian model averaging* (left) uses the unit-information g-prior with the dilution model prior, which penalizes collinearity by down-weighting models whose regressor correlation matrix is close to singular; P. mean, P. SD and PIP are the posterior mean, posterior standard deviation and posterior inclusion probability. *OLS, two-way clustered* (middle) regresses the elasticity on the same moderators with standard errors clustered simultaneously by study and by underlying database. *FMA* (right) is a frequentist model average with smoothed-AIC weights over the models visited by the sampler, reporting each moderator's model-averaged coefficient (*Coef.*) and its smoothed-AIC inclusion weight (*Wt.*, the frequentist analogue of the posterior inclusion probability); because those weights depend on which models the sampler visits over a very large model space, this column is a Monte Carlo estimate that is not exactly reproducible run to run (up to about 0.05), unlike the reproducible Bayesian and OLS columns. The standard error of the estimate absorbs residual publication selection. Scale-laden moderators enter in logs; reference categories are omitted. Variables are defined in Table S10.

Table S14: Bayesian model averaging of the short-run price elasticity

	Bayesian model averaging (UIP g-prior, dilution)			OLS, two-way clustered (study × database)			FMA (smoothed AIC)	
	P. mean	P. SD	PIP	Coef.	SE	p	Coef.	Wt.
Constant	-0.449	–	1.000	-0.766	0.367	0.037	-0.460	1.000
Standard error	-0.846	0.045	1.000	-0.860	0.140	0.000	-0.833	1.000
Estimate type								
Converted from Hicksian	0.033	0.038	0.483	0.055	0.034	0.105	0.062	0.960
Identification								
Design-based	0.259	0.040	1.000	0.258	0.095	0.006	0.250	1.000
Instrumented	0.001	0.005	0.020	0.056	0.049	0.255	0.007	0.186
Panel FE / structural	0.066	0.029	0.901	0.085	0.037	0.020	0.065	0.992
Estimate aggregation								
Aggregation: country	-0.000	0.002	0.009	-0.038	0.031	0.215	-0.001	0.050
Aggregation: micro	-0.021	0.033	0.331	-0.025	0.043	0.559	-0.032	0.566
Sector								
Residential demand	-0.015	0.030	0.238	-0.039	0.034	0.246	-0.018	0.351
Industrial demand	-0.014	0.032	0.175	-0.070	0.046	0.129	-0.021	0.350
Data characteristics								
Time-series data	0.006	0.016	0.154	0.029	0.035	0.409	0.010	0.294
Cross-section data	-0.000	0.005	0.012	-0.006	0.071	0.936	-0.000	0.016
Yearly data	-0.019	0.031	0.312	-0.067	0.040	0.092	-0.048	0.854
Data span	0.003	0.010	0.094	0.019	0.024	0.427	0.009	0.350
Study size	-0.001	0.003	0.093	-0.011	0.007	0.138	-0.002	0.235
Price and tariff								
Average price	-0.001	0.004	0.022	-0.035	0.035	0.315	-0.001	0.034
Marginal price	0.000	0.003	0.012	0.006	0.059	0.921	0.000	0.019
Increasing-block tariff	-0.001	0.005	0.019	-0.035	0.033	0.288	-0.001	0.026
Decreasing-block tariff	0.000	0.003	0.009	0.018	0.049	0.716	0.000	0.011
Time-of-use tariff	0.065	0.045	0.733	0.057	0.063	0.367	0.066	0.858
Demand controls								
Control: demographics	-0.064	0.023	0.948	-0.052	0.034	0.122	-0.054	0.981
Control: temperature	-0.000	0.002	0.010	-0.014	0.027	0.589	-0.001	0.028
Control: appliance stock	-0.083	0.029	0.950	-0.068	0.043	0.111	-0.085	1.000
Control: other fuels	-0.010	0.021	0.232	-0.019	0.032	0.566	-0.027	0.675
Model specification								
Reduced form	0.059	0.025	0.912	0.084	0.030	0.005	0.074	1.000
Static model	-0.133	0.023	1.000	-0.123	0.047	0.009	-0.133	1.000
ARDL	0.000	0.006	0.012	0.015	0.038	0.701	0.001	0.019
Lagged endogenous	-0.001	0.006	0.023	-0.024	0.041	0.551	-0.000	0.021
Linear demand	0.056	0.049	0.627	0.145	0.098	0.136	0.093	0.989
Double-log demand	0.001	0.011	0.030	0.056	0.065	0.390	0.003	0.055
Geography and context								
United States	-0.000	0.002	0.011	-0.005	0.044	0.908	-0.000	0.012
Europe	-0.000	0.003	0.011	-0.026	0.047	0.577	-0.000	0.016
Daylight hours	0.020	0.005	0.996	0.024	0.006	0.000	0.020	1.000
Population	0.000	0.001	0.008	0.009	0.011	0.424	0.000	0.016
Publication characteristics								
Publication year	0.000	0.005	0.015	0.030	0.084	0.723	0.001	0.029
Impact factor	-0.000	0.001	0.007	-0.009	0.021	0.677	-0.000	0.012
Citations	0.000	0.002	0.015	0.010	0.036	0.781	0.001	0.037
Observations	1,647			1,647			1,647	
Studies / databases	226 / 218							

Notes: As in Table S13, estimated on the 1,647 short-run headline estimates; the horizon indicators are constant here and drop out. Variables are defined in Table S11.

Table S15: Bayesian model averaging of the long-run price elasticity

	Bayesian model averaging (UIP g-prior, dilution)			OLS, two-way clustered (study × database)			FMA (smoothed AIC)	
	P. mean	P. SD	PIP	Coef.	SE	p	Coef.	Wt.
Constant	-0.515	–	1.000	-1.525	0.839	0.069	-1.056	1.000
Standard error	-1.072	0.069	1.000	-1.085	0.138	0.000	-1.067	1.000
Estimate type								
Converted from Hicksian	-0.186	0.043	0.995	-0.135	0.067	0.045	-0.202	1.000
Identification								
Design-based	0.001	0.020	0.007	-0.260	0.191	0.174	0.000	0.004
Instrumented	-0.028	0.057	0.221	-0.102	0.077	0.185	-0.050	0.387
Panel FE / structural	0.087	0.071	0.651	0.079	0.082	0.337	0.082	0.629
Estimate aggregation								
Aggregation: country	0.000	0.004	0.008	0.022	0.063	0.729	0.000	0.008
Aggregation: micro	-0.000	0.006	0.008	0.068	0.105	0.518	0.000	0.003
Sector								
Residential demand	-0.000	0.005	0.009	0.044	0.066	0.499	-0.000	0.005
Industrial demand	0.012	0.037	0.113	0.114	0.076	0.133	0.041	0.433
Data characteristics								
Time-series data	0.000	0.004	0.007	-0.002	0.075	0.983	0.000	0.007
Cross-section data	-0.085	0.120	0.376	-0.064	0.114	0.572	-0.042	0.240
Yearly data	-0.005	0.023	0.047	-0.071	0.100	0.477	-0.012	0.142
Data span	0.047	0.043	0.582	0.062	0.046	0.178	0.077	0.910
Study size	0.000	0.001	0.009	-0.012	0.019	0.521	-0.001	0.066
Price and tariff								
Average price	0.000	0.004	0.007	0.099	0.110	0.366	0.000	0.006
Marginal price	0.001	0.011	0.012	0.247	0.173	0.154	0.021	0.160
Increasing-block tariff	0.000	0.005	0.008	-0.054	0.084	0.519	0.000	0.006
Decreasing-block tariff	-0.002	0.018	0.016	-0.028	0.149	0.850	-0.001	0.012
Time-of-use tariff	-0.001	0.016	0.011	0.010	0.183	0.957	0.000	0.007
Demand controls								
Control: demographics	-0.001	0.012	0.021	-0.056	0.075	0.455	-0.003	0.052
Control: temperature	0.000	0.003	0.007	-0.084	0.059	0.156	-0.000	0.009
Control: appliance stock	-0.174	0.073	0.912	-0.151	0.059	0.011	-0.157	0.995
Control: other fuels	-0.000	0.005	0.009	-0.051	0.068	0.457	-0.000	0.011
Model specification								
Reduced form	0.000	0.006	0.010	0.071	0.080	0.373	0.001	0.013
Static model	-0.001	0.008	0.013	-0.069	0.090	0.447	-0.000	0.006
ARDL	0.007	0.032	0.064	0.110	0.066	0.096	0.070	0.536
Lagged endogenous	0.015	0.055	0.086	0.235	0.085	0.006	0.208	0.939
Linear demand	-0.016	0.057	0.088	-0.446	0.179	0.013	-0.158	0.804
Double-log demand	0.001	0.013	0.015	-0.200	0.100	0.047	0.001	0.017
Geography and context								
United States	0.035	0.062	0.270	0.313	0.096	0.001	0.169	0.997
Europe	0.085	0.074	0.632	0.206	0.078	0.008	0.177	1.000
Daylight hours	0.000	0.001	0.011	-0.012	0.016	0.454	-0.000	0.015
Population	0.000	0.001	0.008	-0.007	0.020	0.704	-0.000	0.028
Publication characteristics								
Publication year	0.013	0.053	0.067	0.342	0.197	0.083	0.110	0.451
Impact factor	0.000	0.005	0.013	-0.016	0.048	0.745	0.000	0.004
Citations	0.015	0.030	0.232	0.065	0.033	0.048	0.028	0.525
Observations	723			723			723	
Studies / databases	151 / 143							

Notes: As in Table S13, estimated on the 723 long-run headline estimates; the horizon indicators are constant here and drop out. Variables are defined in Table S12.

Table S16: Model averaging under alternative priors (integrated sample)

Variable	A		B		C		D	
	P. mean	PIP	P. mean	PIP	P. mean	PIP	P. mean	PIP
Standard error	-0.830	1.000	-0.830	1.000	-0.830	1.000	-0.829	1.000
Short run	0.114	1.000	0.114	1.000	0.114	1.000	0.114	1.000
Long run	-0.103	1.000	-0.102	1.000	-0.103	1.000	-0.102	1.000
Cross-section data	-0.166	1.000	-0.166	1.000	-0.166	1.000	-0.161	1.000
Yearly data	-0.125	1.000	-0.125	1.000	-0.125	1.000	-0.124	1.000
Decreasing-block tariff	-0.174	1.000	-0.174	1.000	-0.174	1.000	-0.174	1.000
Lagged endogenous	0.101	1.000	0.100	1.000	0.101	1.000	0.098	1.000
Control: other fuels	-0.074	1.000	-0.074	1.000	-0.074	1.000	-0.073	1.000
Control: demographics	-0.067	0.998	-0.067	0.998	-0.067	0.998	-0.063	1.000
Study size	-0.017	0.995	-0.017	0.997	-0.017	0.995	-0.017	1.000
Design-based	0.159	0.994	0.159	0.994	0.159	0.994	0.154	0.996
Daylight hours	0.019	0.992	0.019	0.993	0.019	0.992	0.019	0.997
Panel FE / structural	0.075	0.989	0.075	0.990	0.075	0.989	0.074	0.998
Impact factor	0.034	0.665	0.034	0.668	0.034	0.665	0.032	0.689
United States	0.037	0.664	0.038	0.670	0.037	0.664	0.036	0.647
Citations	0.018	0.474	0.019	0.488	0.018	0.474	0.019	0.556
Aggregation: micro	-0.022	0.376	-0.022	0.386	-0.022	0.376	-0.023	0.433
Reduced form	0.014	0.338	0.015	0.366	0.014	0.338	0.020	0.509
Population	0.004	0.313	0.005	0.331	0.004	0.313	0.007	0.474
Residential demand	-0.012	0.293	-0.013	0.315	-0.012	0.293	-0.016	0.420
Data span	0.005	0.200	0.006	0.220	0.005	0.201	0.010	0.385
Linear demand	0.010	0.188	0.011	0.207	0.010	0.188	0.021	0.379
Marginal price	0.005	0.118	0.006	0.138	0.005	0.118	0.021	0.414
Control: appliance stock	-0.005	0.118	-0.005	0.130	-0.005	0.118	-0.012	0.268
Double-log demand	-0.004	0.093	-0.004	0.104	-0.004	0.094	-0.005	0.136
Time-series data	0.003	0.074	0.003	0.081	0.003	0.074	0.004	0.118
Europe	0.002	0.046	0.002	0.052	0.002	0.046	0.005	0.117
ARDL	0.002	0.034	0.002	0.039	0.002	0.034	0.003	0.063
Publication year	0.001	0.021	0.001	0.025	0.001	0.022	0.003	0.062
Instrumented	-0.001	0.019	-0.001	0.020	-0.001	0.019	-0.000	0.015
Industrial demand	0.000	0.018	0.000	0.020	0.000	0.018	0.000	0.024
Static model	-0.000	0.016	-0.000	0.018	-0.000	0.016	-0.001	0.043
Time-of-use tariff	0.000	0.014	0.000	0.016	0.000	0.014	0.001	0.026
Converted from Hicksian	-0.000	0.013	-0.000	0.015	-0.000	0.013	-0.000	0.019
Average price	-0.000	0.012	-0.000	0.015	-0.000	0.013	-0.000	0.021
Control: temperature	-0.000	0.010	-0.000	0.012	-0.000	0.010	-0.000	0.012
Aggregation: country	0.000	0.010	0.000	0.011	0.000	0.010	0.000	0.010
Increasing-block tariff	0.000	0.009	0.000	0.010	0.000	0.009	0.000	0.011

Notes: Posterior mean and inclusion probability (PIP) for the integrated model average under four prior specifications, ordered by baseline PIP. (A) unit-information g-prior, dilution model prior (baseline); (B) unit-information g-prior, uniform model prior; (C) BRIC g-prior, random (beta-binomial) model prior; (D) Hannan–Quinn g-prior, dilution model prior. The key moderators keep their sign and inclusion across priors.

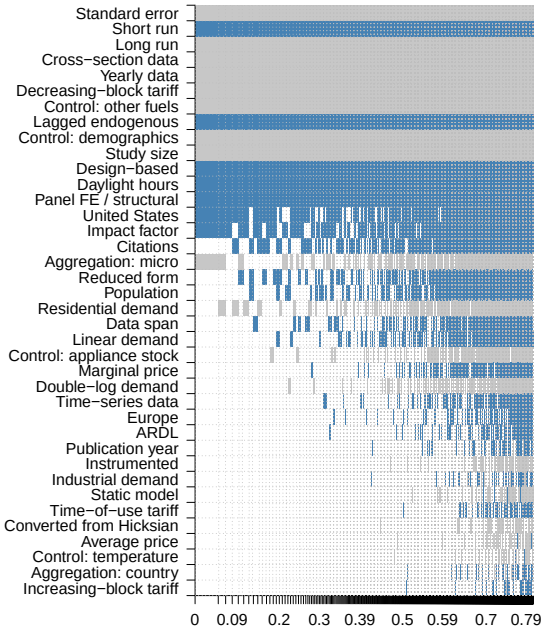
Table S17: Model-averaging diagnostics

Model	Obs.	E[model size]	Models visited	Corr PMP	Shrinkage
Integrated	3,324	17.11	146,615	1.00	1.00
Short run	1,647	11.43	148,177	1.00	1.00
Long run	723	6.54	117,429	1.00	1.00

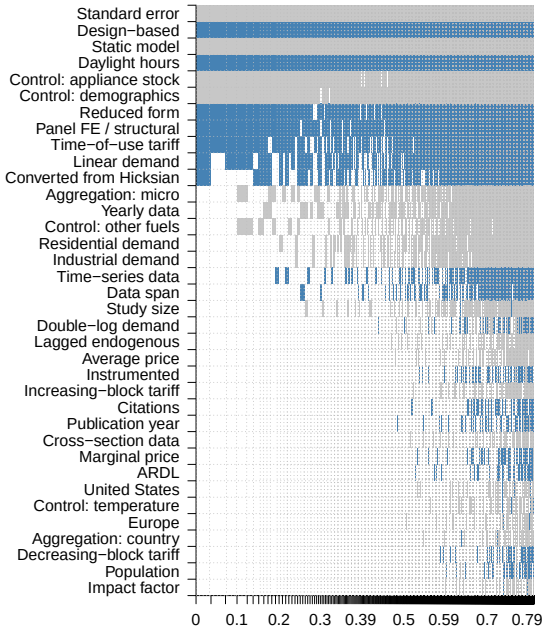
Notes: Birth–death MCMC (1.5×10^5 draws after 5×10^4 burn-in) for each averaging. "E[model size]" is the sum of inclusion probabilities; "Corr PMP" is the correlation between analytical and MCMC posterior model probabilities (one indicates convergence); "Shrinkage" is $g/(1+g)$.

What the averaging shows. Four features reinforce the body. First, publication selection is the most robust regularity in the corpus: the standard-error term enters every model in every subsample (PIP = 1.00) with a negative posterior mean that grows with the horizon (-0.83 integrated, -0.85 short run, -1.07 long run), the funnel asymmetry of Table 2 recovered as a model-averaged coefficient that survives against every moderator at once. Second, and most telling for the flexibility premise, the best-identified and most technology-rich settings are, conditional on everything else, the *least* price-responsive: in the short-run averaging the design-based tier ($+0.26$, PIP = 1.00, and OLS-significant) is robustly included and pulls the elasticity toward zero, while the time-of-use indicator ($+0.07$, PIP = 0.73, its OLS coefficient not distinguishable from zero) is weakly included and only sign-consistent, the model-averaging counterpart of the identification ladder and of the new-regime and time-of-use findings (Figure 4, Section 6). Third, the vintage dimension itself carries *negligible* weight: in the integrated model the publication-year moderator is almost never selected and the data span only rarely (PIP = 0.02 and 0.20), both entering with posterior means near zero, so once composition is averaged over there is no heterogeneity along the calendar axis a rising-flexibility world would light up, the averaging counterpart of the flat *publication* clock; the data mid-year itself is not among the model-averaging moderators, so this appendix speaks to the reporting clock rather than directly to the behavioral one. Fourth, the heterogeneity that *is* robust is compositional and methodological: data structure and frequency (cross-section -0.17, yearly -0.12, both PIP = 1.00), study size, the dynamic-adjustment indicators (lagged endogenous $+0.10$ integrated, static model -0.13 short run), decreasing-block tariffs (-0.17) and a demographics control, exactly the study characteristics the ladder and horizon analyses hold fixed. The frequentist model average and the two-way clustered OLS agree in sign with the posterior means on every reliably included moderator, the core integrated-sample terms are stable across the four prior specifications (Table S16), while the horizon-specific positives (the short-run design-based and time-of-use coefficients) are checked only for sign in that integrated sweep, and the chains have converged (posterior-model-probability correlation 1.00; Table S17), with the regressor correlations (Figure S8) justifying the dilution prior.

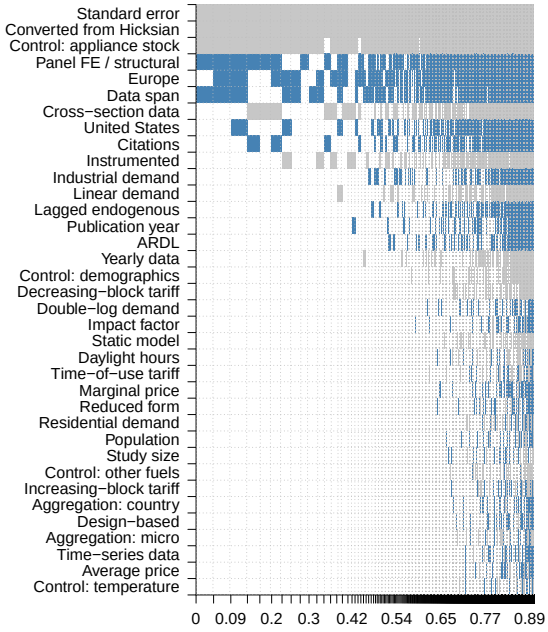
Figure S7: Model inclusion in Bayesian model averaging



(a) Full headline sample



(b) Short-run estimates



(c) Long-run estimates

Notes: Each column is a regression model; column width is proportional to the model's posterior probability, cumulated along the horizontal axis. Rows are regressors, sorted by posterior inclusion probability. Blue means the variable enters with a positive coefficient, grey negative, white excluded. The estimation uses the unit-information g-prior and the dilution model prior.

Figure S8: Regressor correlation matrix (collinearity diagnostic)

Notes: Pairwise correlations among the integrated-model regressors; cells with $|r| < 0.15$ are left blank and $|r| \geq 0.4$ are bold. Blue = positive, grey = negative. The visible clusters (among the identification, model-form and data-structure indicators) are why we average under the collinearity-penalizing dilution prior rather than reporting a single regression.

S6.1 Best-practice prediction, country by country

The model average also lets us read off a *best-practice* elasticity for each country: the response an ideally-designed study would report there. We predict from the integrated two-way-clustered regression of Table S13, evaluated not at the average study but at a best-practice profile $x^\star$. Following the variable list, (i) the standard error is set to zero, stripping out the publication-selection component; (ii) the journal impact factor and citation count are set to their 95th percentiles, a well-cited outlet without letting a single citation outlier drive the prediction; (iii) the country-level moderators (the US and Europe indicators, daylight hours and log population) are set to *each country's own level*; (iv) the identification tier is set to best practice, with the design-based indicator on and the instrumented and panel-FE/structural indicators off (naive is the reference); (v) the demand equation is fully controlled (demographics, temperature, appliance stock, and substitute-fuel prices) and estimated on micro, household-level data; and (vi) every remaining moderator (model specification and functional form) is held at its sample mean, so the prediction describes a fully-specified, design-based, de-biased study conducted in that country. We report the prediction at both horizons: the short-run profile sets the short-run indicator on, the long-run profile the long-run indicator, with the other horizon dummies off.

The best-practice elasticity is the linear prediction $\text{Mean}^\star = x^{\star\prime}\widehat{\beta}$, and because it is a linear combination of the estimated coefficients its sampling variance is $x^{\star\prime}\widehat{V}\,x^\star$, where $\widehat{V}$ is the two-way (study $\times$ database) clustered covariance; we report the Wald interval $\text{Mean}^\star \pm 1.96\sqrt{x^{\star\prime}\widehat{V}\,x^\star}$. Figure S9 shows the two predictions for the countries with at least thirty headline estimates, ranked by the short-run value; Table S18 lists every country with at least five. The message mirrors the pooled result: the best-practice short-run elasticity is small and negative everywhere (typically around -0.18), in line with the corrected short-run consensus of the body, though its confidence interval still spans zero in fewer than half the countries; the long-run prediction is roughly twice as large and significantly negative in most countries, the horizon gradient reproducing country by country. Because only the four country-level moderators vary across predictions while all method, sector, and horizon slopes are held common, the compressed cross-country spread is imposed by the pooled model rather than demonstrated as genuine homogeneity. The cross-country spread is modest relative to the within-country uncertainty, so the audit's central finding travels: no country's best-practice *prediction* supports the large, rising responsiveness the flexibility premise assumes. For countries without an in-country design-based study, this value is a model extrapolation that sets the design-based indicator on, not an observed record.

Table S18: Best-practice price elasticity by country

	Short run		Long run	
Country	Mean*	95% CI	Mean*	95% CI
Argentina	−0.201	[−0.374, −0.028]	−0.419	[−0.588, −0.250]
Australia	−0.218	[−0.395, −0.042]	−0.437	[−0.608, −0.266]
Austria	−0.148	[−0.339, 0.044]	−0.366	[−0.546, −0.185]
Azerbaijan	−0.469	[−0.757, −0.182]	−0.687	[−0.970, −0.405]
Bangladesh	−0.189	[−0.365, −0.013]	−0.407	[−0.583, −0.232]
Brazil	−0.206	[−0.386, −0.026]	−0.424	[−0.603, −0.246]
Canada	−0.158	[−0.328, 0.011]	−0.377	[−0.542, −0.212]
Chile	−0.215	[−0.390, −0.039]	−0.433	[−0.603, −0.263]
China	−0.145	[−0.327, 0.038]	−0.363	[−0.548, −0.178]
Colombia	−0.235	[−0.416, −0.054]	−0.453	[−0.630, −0.277]
Costa Rica	−0.265	[−0.455, −0.076]	−0.483	[−0.664, −0.302]
Denmark	−0.124	[−0.313, 0.066]	−0.342	[−0.520, −0.164]

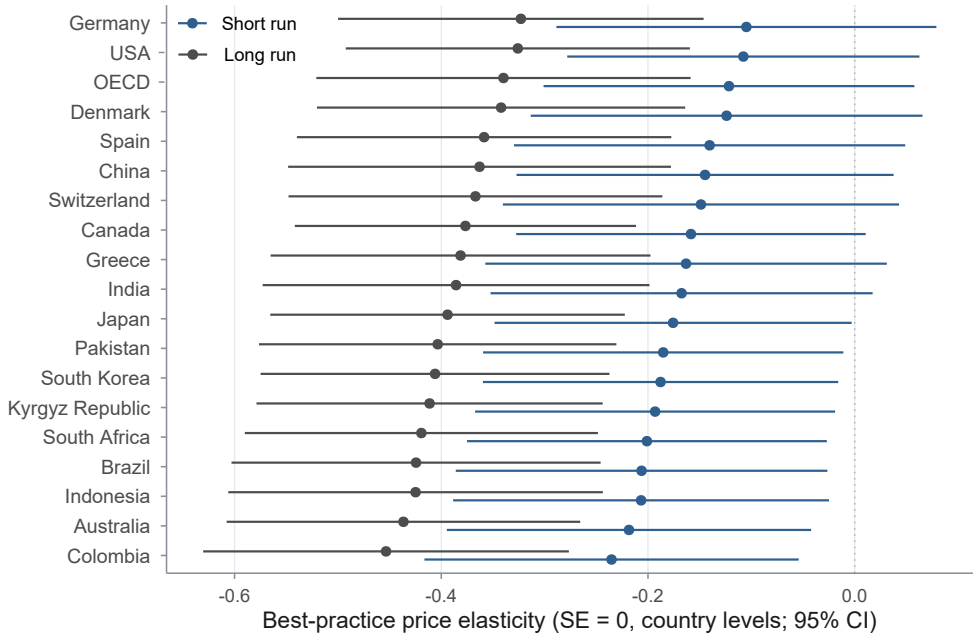
Continued on next page

Table S18 continued

	Short run		Long run	
Country	Mean*	95% CI	Mean*	95% CI
Egypt	−0.198	[−0.372, −0.024]	−0.416	[−0.588, −0.245]
EU	−0.091	[−0.278, 0.096]	−0.309	[−0.493, −0.126]
Finland	−0.064	[−0.254, 0.125]	−0.282	[−0.461, −0.104]
France	−0.123	[−0.309, 0.063]	−0.341	[−0.520, −0.162]
Germany	−0.105	[−0.289, 0.079]	−0.323	[−0.500, −0.146]
Ghana	−0.237	[−0.418, −0.056]	−0.455	[−0.631, −0.280]
Greece	−0.163	[−0.357, 0.031]	−0.381	[−0.565, −0.198]
Hong Kong	−0.236	[−0.416, −0.056]	−0.454	[−0.628, −0.280]
India	−0.167	[−0.352, 0.017]	−0.386	[−0.573, −0.199]
Indonesia	−0.207	[−0.389, −0.025]	−0.425	[−0.606, −0.244]
Iran	−0.188	[−0.362, −0.015]	−0.406	[−0.577, −0.236]
Ireland	−0.141	[−0.334, 0.052]	−0.359	[−0.540, −0.178]
Israel	−0.226	[−0.404, −0.047]	−0.444	[−0.616, −0.272]
Italy	−0.131	[−0.318, 0.057]	−0.349	[−0.529, −0.169]
Jamaica	−0.255	[−0.442, −0.068]	−0.473	[−0.652, −0.295]
Japan	−0.176	[−0.349, −0.003]	−0.394	[−0.565, −0.222]
Kazakhstan	−0.180	[−0.351, −0.009]	−0.398	[−0.564, −0.232]
Kuwait	−0.245	[−0.431, −0.059]	−0.463	[−0.641, −0.286]
Kyrgyz Republic	−0.193	[−0.367, −0.019]	−0.411	[−0.579, −0.244]
Malaysia	−0.242	[−0.424, −0.060]	−0.460	[−0.637, −0.283]
Mexico	−0.200	[−0.376, −0.025]	−0.418	[−0.592, −0.245]
Middle income	−0.076	[−0.274, 0.123]	−0.294	[−0.492, −0.095]
Mozambique	−0.458	[−0.743, −0.172]	−0.676	[−0.957, −0.395]
Netherlands	−0.122	[−0.308, 0.064]	−0.340	[−0.517, −0.164]
New Zealand	−0.218	[−0.397, −0.039]	−0.436	[−0.607, −0.265]
Niger	−0.247	[−0.430, −0.063]	−0.465	[−0.641, −0.289]
Nigeria	−0.217	[−0.396, −0.038]	−0.435	[−0.611, −0.259]
Norway	−0.089	[−0.278, 0.100]	−0.307	[−0.485, −0.130]
OECD	−0.122	[−0.301, 0.058]	−0.340	[−0.521, −0.159]
Pakistan	−0.185	[−0.360, −0.011]	−0.403	[−0.576, −0.231]
Peru	−0.231	[−0.411, −0.052]	−0.449	[−0.625, −0.274]
Philippines	−0.216	[−0.395, −0.038]	−0.435	[−0.610, −0.259]
Portugal	−0.161	[−0.355, 0.032]	−0.380	[−0.563, −0.196]
Saudi Arabia	−0.222	[−0.399, −0.045]	−0.441	[−0.612, −0.269]
South Africa	−0.201	[−0.375, −0.027]	−0.419	[−0.590, −0.248]
South Korea	−0.188	[−0.360, −0.016]	−0.406	[−0.575, −0.237]
Spain	−0.140	[−0.330, 0.049]	−0.359	[−0.540, −0.177]
Sri Lanka	−0.239	[−0.420, −0.058]	−0.457	[−0.633, −0.281]
Sweden	−0.078	[−0.263, 0.107]	−0.296	[−0.471, −0.121]
Switzerland	−0.149	[−0.340, 0.043]	−0.367	[−0.548, −0.186]
Taiwan	−0.217	[−0.394, −0.041]	−0.436	[−0.607, −0.264]
Tunisia	−0.215	[−0.391, −0.039]	−0.433	[−0.603, −0.263]
Turkey	−0.138	[−0.328, 0.051]	−0.357	[−0.538, −0.175]
Ukraine	−0.150	[−0.342, 0.041]	−0.369	[−0.552, −0.185]
United Kingdom	−0.099	[−0.282, 0.084]	−0.317	[−0.493, −0.142]
USA	−0.108	[−0.278, 0.063]	−0.326	[−0.492, −0.159]
Yemen	−0.237	[−0.417, −0.056]	−0.455	[−0.629, −0.280]

Notes: Mean* is the best-practice predicted elasticity, the fitted value at a country's best-practice profile $x^\star$; the table reports short- and long-run Mean* with 95% two-way (study × database) clustered intervals, for the countries contributing at least five headline estimates, ordered alphabetically; a few rows are country aggregates (EU, OECD, Middle income) rather than individual countries. Country-level moderators are set to each country's own level; all other moderators follow the best-practice profile of Figure S9.

Figure S9: Best-practice price elasticity by country, short and long run



Notes: Best-practice prediction from the integrated regression of Table S13 (standard error set to zero, impact factor and citations at their 95th percentiles, country-level moderators at each country's own level, identification at best practice (design-based tier on; instrumented and structural tiers off, naive reference), demand controls on and micro-level data, remaining moderators at sample means), with 95% intervals from the two-way (study × database) clustered covariance. Countries with at least thirty headline estimates, ranked by the short-run prediction.

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