

Weighted Adjustment Instrumental Variable Estimator

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Reported precision is a choice

Table: Regression of test scores on class size

	Krueger (1999)	Replications using different computations of SE				
	(1)	(2) Bootstrap of class clusters	(3) Class clusters	(4) School clusters	(5) Huber-White SE	(6) Plain vanilla SE
Small class	4.82 (2.19)	4.82 (2.00)	4.82 (1.88)	4.82 (1.38)	4.82 (0.79)	4.82 (0.76)
Sample	5,861	5,861	5,861	5,861	5,861	5,861

Notes: Dependent variable: test score percentile. Standard errors (SE) in parentheses.

Meta-analysis of the class size effect

Available at meta-analysis.cz/class. Spoiler: effect size tiny.

Changing controls changes precision

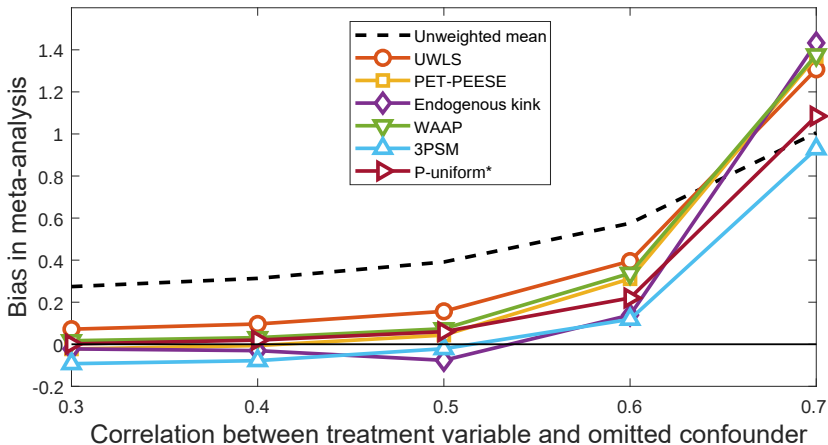
Table: Regression of test scores on class size

	(1)	(2)	(3)	(4)
Small class (treatment)	4.82 (2.19)	5.37 (1.26)	5.36 (1.21)	5.37 (1.19)
White/Asian			8.35 (1.35)	8.44 (1.36)
Girl			4.48 (0.63)	4.39 (0.63)
Free lunch			-13.15 (0.77)	-13.07 (0.77)
White teacher				-0.57 (2.1)
Teacher experience				0.26 (0.10)
Master's degree				-0.51 (1.06)
School intercepts	No	Yes	Yes	Yes
Sample	5,861	5,861	5,861	5,861

Notes: Adapted from Krueger (1999). Dependent variable: test score percentile.

Standard errors in parentheses.

Cures worse than the disease?



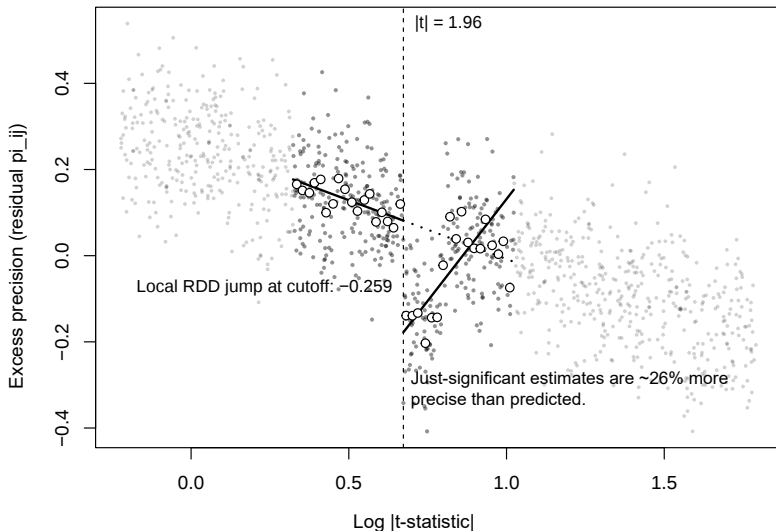
Diagnostic: excess precision

Residual from meta-regression:

$$\log \text{SE}(\hat{E}_i) = \alpha_0 + \alpha_1 \log N_i + \underbrace{\pi_i}_{<0: \text{excess precision}}$$

- Negative $\pi_i \rightarrow$ reported **SE too small relative to N** $\rightarrow$ may indicate p -hacking.
- Interpretation: $\pi_i = -0.25$ means SE is roughly one quarter smaller than expected.

Residual discontinuity test



Weighted Adjustment Instrumental Variable Estimator

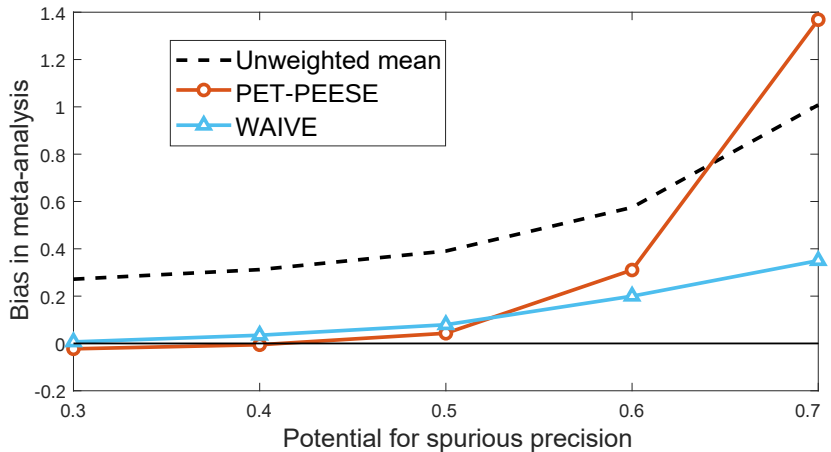
WAIVE combines two adjustments:

- 1 Replace reported SEs by the component predicted from sample size (instrument).
- 2 Downweight spuriously precise estimates **near significance cutoffs**. Exponential tilt:

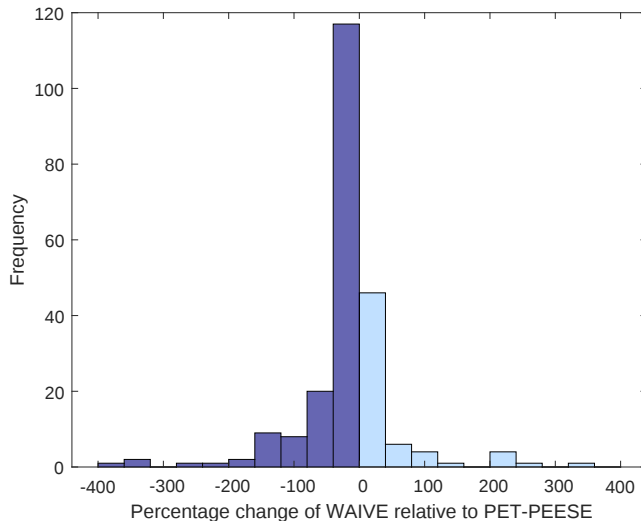
$$\omega_i = \exp[-D_i \max(0, -\pi_i)],$$

- D_i is high when $|t_i|$ is close to 1.96.
- Weight $\omega_i \in (0, 1]$; only excess precision penalized.

WAIVE alleviates the bias in simulations



WAIVE shrinks PET-PEESE in applications



Practical takeaway

Bottom line

MAIVE (*Nature Comms* 2025) instruments the standard error;
WAIVE adds a downweighting layer for studies with **excess precision near significance cutoffs**.

When to use

- Can methods or p -hacking influence both $\hat{E}$ and SE ?
- Yes → MAIVE / WAIVE help (**EasyMeta.org**).
- No → MAIVE / WAIVE may be too conservative; standard methods suffice.

Project website

meta-analysis.cz/maive

EasyMeta.org! Papers on instrumental meta:



Irsova Z., P. Bom, T. Havranek, & H. Rachinger (2025):
Spurious Precision in Meta-Analysis of Observational
Research.

Nature Communications 16: 8454.



Irsova Z., T. Havranek, X. Bortnikova, & F. Bartos (2026):
Meta-Analysis of Field Studies on Beauty and Professional
Success.

Nature Human Behaviour, R&R.



Cala P., T. Havranek, Z. Irsova, M. Luskova, J. Matousek
& J. Novak (2026): Financial Incentives and Performance:
A Meta-Analysis of Experiments in Economics.
Journal of Political Economy Microeconomics, forthcoming.

Backup slides for discussion.

Where to look for more

- **Method paper.** Irsova, Bom, Havranek & Rachinger (2025), *Nature Communications* 16:8454.
- **Tools and replication.** easymeta.org · meta-analysis.cz/maive · MAIVE package on CRAN

Excess precision: log first-stage residual

First stage (log form; extension over published MAIVE)

$$\log(\text{SE}_i) = \alpha + \beta \log(N_i) + \gamma' X_i + \pi_i.$$

- $\pi_i < 0 \Rightarrow$ reported SE **smaller than sample size predicts**.
- X_i : pre-specified design features; *not* researcher-chosen inference tweaks.
- Interpretation: $\pi_i = -0.10 \Rightarrow \text{SE} \approx 10\%$ below prediction (exact: $1 - e^{-0.10} = 9.5\%$).

Foundational first stage from MAIVE (Nature Comms 2025); log version used going forward.

Downweighting: D_i on $\log |t|$

Triangular danger-zone kernel (Chania operational default)

$$D_i = \max\left\{0, 1 - \frac{|\log |t_i| - \log(1.96)|}{h}\right\}, \quad t_i = \hat{\beta}_i/\text{SE}_i.$$

- **Compact support:** $D_i = 0$ outside bandwidth h ; same scale as RDT.

WAIVE weight

$$\omega_i = \exp[-D_i \max(0, -\pi_i)] \in (0, 1].$$

Penalizes only estimates with excess precision *and* $|t|$ near the cutoff.

RDT: residual discontinuity at the significance cutoff

Test of continuity in $E[\pi_i | r_i]$ at $r = c$

$$r_i = \log |t_i|, \quad c = \log(1.96),$$
$$\tau_{\text{RDT}} = \lim_{r \downarrow c} E[\pi_i | r_i = r] - \lim_{r \uparrow c} E[\pi_i | r_i = r].$$

- Null: $E[\pi_i | r_i]$ **continuous** at c .
- $\tau_{\text{RDT}} < 0$: just-significant SEs systematically below prediction.
- $\tau = -0.10$: SEs $\approx 10\%$ smaller than predicted ($1 - e^{-0.10} = 9.5\%$ exact).

RDT: estimation and falsification battery

Estimation

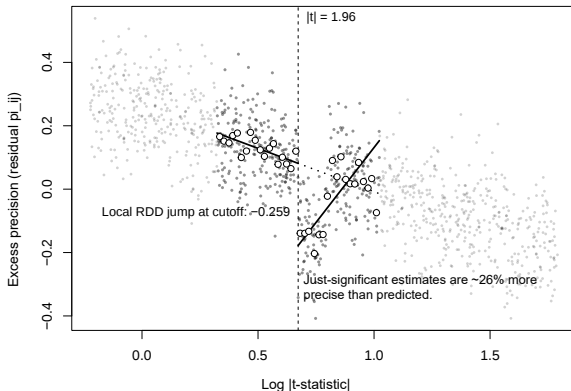
- Outcome: $\hat{\pi}_i$
- Running variable: $\log |t_i|$
- Local linear, both sides
- Triangular kernel
- MSE-optimal bandwidth (Calonico et al. 2014)
- Cluster bootstrap by study

Falsification battery

- Donut hole around $|t| = 1.96$
- Placebo cutoffs ($|t| = 1.5, 2.5$)
- Density tests (McCrary; Cattaneo et al.)
- Bandwidth sensitivity
- Mechanism check: residualize $\log |\hat{\beta}|$ as placebo outcome

No global polynomial; identification is intentionally local.

RDT: the picture



Local jump $\hat{\tau}_{\text{RDT}} \approx -0.259$: substantial concentration of **excess precision** just above $|t| = 1.96$.

Simulation: data-generating process

Primary studies

$$Y = \alpha_0 + \alpha_1 X_1 + \alpha_2 X_2 + u, \quad X_2 = \psi X_1 + (1 - \psi) \varepsilon.$$

- True effect $\alpha_1 = 1$ (bias on the same scale).
- X_2 : true control; replacing it with weaker proxies creates OVB, **stronger as ψ rises**.
- $M = 80$ studies; sample sizes calibrated to published meta-analyses.
- 50% potential p -hackers; search over proxy controls until $t > 1.96$ or $H = 50$ attempts.

Simulation design: MAIVE (Nature Comms 2025); tools at easymeta.org.

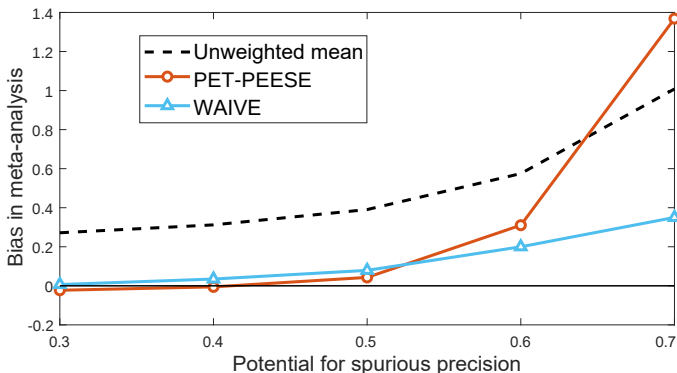
Simulation: what is being selected?

Specification search moves *both* estimate and precision

Changing the conditioning set perturbs $\hat{\beta}_i$ and SE_i simultaneously; the accepted specification may have a larger estimate, a smaller SE, or both.

- **Estimate channel.** Accepted spec yields larger $|\hat{\beta}_i|$.
- **Precision channel.** Accepted spec yields smaller SE_i — the MAIVE/WAIVE target.
- Conventional funnel-based corrections treat SE_i as exogenous and cannot see the precision channel.

Simulation: WAIVE alleviates the bias



WAIVE bias stays close to zero as the potential for spurious precision rises;
PET-PEESE diverges.

Experimental WAIVE layer; simulation design from MAIVE (Nature Comms 2025).